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B. P. Abbott

California Institute of Technology

K. AultONeal

Embry-Riddle Aeronautical University

S. Gaudio

Embry-Riddle Aeronautical University

K. Gill

Embry-Riddle Aeronautical University

B. Hughey

Embry-Riddle Aeronautical University

See next page for additional authors

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Authors

B. P. Abbott, K. AultONeal, S. Gaudio, K. Gill, B. Hughey, J. W. W. Pratt, E. Schmidt, S. G. Schwalbe, M. J. Szczepańczyk, M. Zanolin, and et al.



Upper Limits on Gravitational Waves from Scorpius X-1 from a Model-based Cross-correlation Search in Advanced LIGO Data

B. P. Abbott¹, R. Abbott¹, T. D. Abbott², F. Acernese^{3,4}, K. Ackley⁵, C. Adams⁶, T. Adams⁷, P. Addesso⁸, R. X. Adhikari¹, V. B. Adya⁹, C. Affeldt⁹, M. Afrough¹⁰, B. Agarwal¹¹, M. Agathos¹², K. Agatsuma¹³, N. Aggarwal¹⁴, O. D. Aguiar¹⁵, L. Aiello^{16,17}, A. Ain¹⁸, P. Ajith¹⁹, B. Allen^{9,20,21}, G. Allen¹¹, A. Allocca^{22,23}, P. A. Altin²⁴, A. Amato²⁵, A. Ananyeva¹, S. B. Anderson¹, W. G. Anderson²⁰, S. Antier²⁶, S. Appert¹, K. Arai¹, M. C. Araya¹, J. S. Areeda²⁷, N. Arnaud^{26,28}, K. G. Arun²⁹, S. Ascenzi^{17,30}, G. Ashton⁹, M. Ast³¹, S. M. Aston⁶, P. Astone³², P. Aufmuth²¹, C. Aubert⁹, K. AultO'Neal³³, A. Avila-Alvarez²⁷, S. Babak³⁴, P. Bacon³⁵, M. K. M. Bader¹³, S. Bae³⁶, P. T. Baker^{37,38}, F. Baldaccini^{39,40}, G. Ballardin²⁸, S. W. Ballmer⁴¹, S. Banagiri⁴², J. C. Barayoga¹, S. E. Barclay⁴³, B. C. Barish¹, D. Barker⁴⁴, F. Barone^{3,4}, B. Barr⁴³, L. Barsotti¹⁴, M. Barsuglia³⁵, D. Barta⁴⁵, J. Bartlett⁴⁴, I. Bartos⁴⁶, R. Bassiri⁴⁷, A. Basti^{22,23}, J. C. Batch⁴⁴, C. Baune⁹, M. Bawaj^{48,40}, M. Bazzan^{49,50}, B. Bécsy⁵¹, C. Beer⁹, M. Bejger⁵², I. Belahcene²⁶, A. S. Bell⁴³, B. K. Berger¹, G. Bergmann⁹, C. P. L. Berry⁵³, D. Bersanetti^{54,55}, A. Bertolini¹³, J. Betzwieser¹⁴, S. Bhagwat⁵⁶, R. Bhandare⁵⁶, I. A. Bilenko⁵⁷, G. Billingsley¹, C. R. Billman⁵⁸, J. Birch⁶, R. Birney⁵⁸, O. Birnholtz⁹, S. Biscans¹⁴, A. Bisht²¹, M. Bitossi^{23,28}, C. Biwer⁴¹, M. A. Bizouard²⁶, J. K. Blackburn¹, J. Blackman⁵⁹, C. D. Blair⁶⁰, D. G. Blair⁶⁰, R. M. Blair⁴⁴, S. Bloemen⁶¹, O. Bock⁹, N. Bode⁹, M. Boer⁶², G. Bogaert⁶², A. Bohe³⁴, F. Bondu⁶³, R. Bonnand⁷, B. A. Boom¹³, R. Bork¹, V. Boschi^{22,23}, S. Bose^{18,64}, Y. Bouffanais³⁵, A. Bozzi²⁸, C. Bradaschia²³, P. R. Brady²⁰, V. B. Braginsky^{57,149}, M. Branchesi^{65,66}, J. E. Brau⁶⁷, T. Briant⁶⁸, A. Brillet⁶², M. Brinkmann⁹, V. Brisson²⁶, P. Brockill²⁰, J. E. Broida⁶⁹, A. F. Brooks¹, D. A. Brown⁴¹, D. D. Brown⁵³, N. M. Brown¹⁴, S. Brunett¹, C. C. Buchanan², A. Buikema¹⁴, T. Bulik⁷⁰, H. J. Bulten^{13,71}, A. Buonanno^{34,72}, D. Buskulic⁷, C. Buy³⁵, R. L. Byer⁴⁷, M. Cabero⁹, L. Cadonati⁷³, G. Cagnoli^{25,74}, C. Cahillane¹, J. Calderón Bustillo⁷³, T. A. Callister¹, E. Calloni^{4,75}, J. B. Camp⁷⁶, M. Canepa^{54,55}, P. Canizares⁶¹, K. C. Cannon⁷⁷, H. Cao⁷⁸, J. Cao⁷⁹, C. D. Capano⁹, E. Capocasa³⁵, F. Carbognani²⁸, S. Caride⁸⁰, M. F. Carney⁸¹, J. Casanueva Diaz²⁶, C. Casentini^{17,30}, S. Caudill²⁰, M. Cavaglià¹⁰, F. Cavalier²⁶, R. Cavalieri²⁸, G. Cella²³, C. B. Cepeda¹, L. Cerboni Baiardi^{65,66}, G. Cerretani^{22,23}, E. Cesarini^{17,30}, S. J. Chamberlin⁸², M. Chan⁴³, S. Chao⁸³, P. Charlton⁸⁴, E. Chassande-Mottin³⁵, D. Chatterjee²⁰, K. Chatziioannou⁸⁵, B. D. Cheeseboro^{37,38}, H. Y. Chen⁸⁶, Y. Chen⁵⁹, H.-P. Cheng⁵, A. Chincarini⁵⁵, A. Chiummo²⁸, T. Chmiel⁸¹, H. S. Cho⁸⁷, M. Cho⁷², J. H. Chow²⁴, N. Christensen^{62,69}, Q. Chu⁶⁰, A. J. K. Chua¹², S. Chua⁶⁸, A. K. W. Chung⁸⁸, S. Chung⁶⁰, G. Ciani⁵, R. Ciolfi^{89,90}, C. E. Cirelli⁴⁷, A. Cirone^{54,55}, F. Clara⁴⁴, J. A. Clark⁷³, F. Cleva⁶², C. Cocchieri¹⁰, E. Coccia^{16,17}, P.-F. Cohadon⁶⁸, A. Colla^{32,91}, C. G. Collette⁹², L. R. Cominsky⁹³, M. Constancio Jr.¹⁵, L. Conti⁵⁰, S. J. Cooper⁵³, P. Corban⁶, T. R. Corbitt², K. R. Corley⁴⁶, N. Cornish⁹⁴, A. Corsi⁸⁰, S. Cortese²⁸, C. A. Costa¹⁵, M. W. Coughlin⁶⁹, S. B. Coughlin^{95,96}, J.-P. Coulon⁶², S. T. Countryman⁴⁶, P. Couvares¹, P. B. Covas⁹⁷, E. E. Cowan⁷³, D. M. Coward⁶⁰, M. J. Cowart⁶, D. C. Coyne¹, R. Coyne⁸⁰, J. D. E. Creighton²⁰, T. D. Creighton⁹⁸, J. Cripe², S. G. Crowder⁹⁹, T. J. Cullen²⁷, A. Cumming⁴³, L. Cunningham⁴³, E. Cuoco²⁸, T. Dal Canton⁷⁶, S. L. Danilishin^{9,21}, S. D'Antonio¹⁷, K. Danzmann^{9,21}, A. Dasgupta¹⁰⁰, C. F. Da Silva Costa⁵, V. Dattilo²⁸, I. Dave⁵⁶, M. Davier²⁶, D. Davis⁴¹, E. J. Daw¹⁰¹, B. Day⁷³, S. De⁴¹, D. DeBra⁴⁷, E. Deelman¹⁰², J. Degallaix²⁵, M. De Laurentis^{4,75}, S. Deléglise⁶⁸, W. Del Pozzo^{22,23,53}, T. Denker⁹, T. Dent⁹, V. Dergachev³⁴, R. De Rosa^{4,75}, R. T. De Rosa⁶, R. DeSalvo¹⁰³, J. Devenson⁵⁸, R. C. Devine^{37,38}, S. Dhurandhar¹⁸, M. C. Díaz⁹⁸, L. Di Fiore¹⁴, M. Di Giovanni^{90,104}, T. Di Girolamo^{4,46,75}, A. Di Lieto^{22,23}, S. Di Pace^{32,91}, I. Di Palma^{32,91}, F. Di Renzo^{22,23}, Z. Doctor⁸⁶, V. Dolique²⁵, F. Donovan¹⁴, K. L. Dooley¹⁰, S. Doravari⁹, I. Dorrington⁹⁶, R. Douglas⁴³, M. Dovale Álvarez⁵³, T. P. Downes²⁰, M. Drago⁹, R. W. P. Drever^{1,150}, J. C. Driggers⁴⁴, Z. Du⁷⁹, M. Ducrot⁷, J. Duncan⁹⁵, S. E. Dwyer⁴⁴, T. B. Edo¹⁰¹, M. C. Edwards⁶⁹, A. Effler⁶, H.-B. Eggenstein⁹, P. Ehrens¹, J. Eichholz¹, S. S. Eikenberry⁵, R. A. Eisenstein¹⁴, R. C. Essick¹⁴, Z. B. Etienne^{37,38}, T. Etzel¹, M. Evans¹⁴, T. M. Evans⁶, M. Factourovich⁴⁶, V. Fafone^{16,17,30}, H. Fair⁴¹, S. Fairhurst⁹⁶, X. Fan⁷⁹, S. Farinon⁵⁵, B. Farr⁸⁶, W. M. Farr⁵³, E. J. Fauchon-Jones⁹⁶, M. Favata¹⁰⁵, M. Fays⁹⁶, H. Fehrmann⁹, J. Feicht¹, M. M. Fejer⁴⁷, A. Fernandez-Galiana¹⁴, I. Ferrante^{22,23}, E. C. Ferreira¹⁵, F. Ferrini²⁸, F. Fidecaro^{22,23}, I. Fiori²⁸, D. Fiorucci³⁵, R. P. Fisher⁴¹, R. Flaminio^{25,106}, M. Fletcher⁴³, H. Fong⁸⁵, P. W. F. Forsyth²⁴, S. S. Forsyth⁷³, J.-D. Fournier⁶², S. Frasca^{32,91}, F. Frasconi²³, Z. Frei⁵¹, A. Freise⁵³, R. Frey⁶⁷, V. Frey²⁶, E. M. Fries¹, P. Fritschel¹⁴, V. V. Frolov⁶, P. Fulda^{5,76}, M. Fyffe⁶, H. Gabbard⁹, M. Gabel¹⁰⁷, B. U. Gadre¹⁸, S. M. Gaebel⁵³, J. R. Gair¹⁰⁸, D. K. Galloway¹⁰⁹, L. Gammaitoni³⁹, M. R. Ganija⁷⁸, S. G. Gaonkar¹⁸, F. Garufi^{4,75}, S. Gaudio³³, G. Gaur¹¹⁰, V. Gayathri¹¹¹, N. Gehrels^{76,151}, G. Gemme⁵⁵, E. Genin²⁸, A. Gennai²³, D. George¹¹, J. George⁵⁶, L. Gergely¹¹², V. Germain⁷, S. Ghonge⁷³, Abhirup Ghosh¹⁹, Archisman Ghosh^{13,19}, S. Ghosh^{13,61}, J. A. Giaime^{2,6}, K. D. Giardino⁶, A. Giazotto²³, K. Gill³³, L. Glover¹⁰³, E. Goetz⁹, R. Goetz⁵, S. Gomes⁹⁶, G. González², J. M. Gonzalez Castro^{22,23}, A. Gopakumar¹¹³, M. L. Gorodetsky⁵⁷, S. E. Gossan¹, M. Gosselin²⁸, R. Gouaty⁷, A. Grado^{4,114}, C. Graef⁴³, M. Granata²⁵, A. Grant⁴³, S. Gras¹⁴, C. Gray⁴⁴, G. Greco^{65,66}, A. C. Green⁵³, P. Groot⁶¹, H. Grote⁹, S. Grunewald³⁴, P. Gruning²⁶, G. M. Guidi^{65,66}, X. Guo⁷⁹, A. Gupta⁸², M. K. Gupta¹⁰⁰, K. E. Gushwa¹, E. K. Gustafson¹, R. Gustafson¹¹⁵, B. R. Hall⁶⁴, E. D. Hall¹, G. Hammond⁴³, M. Haney¹¹³, M. M. Hanke⁹, J. Hanks⁴⁴, C. Hanna⁸², O. A. Hannuksela⁸⁸, J. Hanson⁶, T. Hardwick², J. Harms^{65,66}, G. M. Harry¹¹⁶, I. W. Harry³⁴, M. J. Hart⁴³, C.-J. Haster⁸⁵, K. Haughian⁴³, J. Healy¹¹⁷, A. Heidmann⁶⁸, M. C. Heintze⁶, H. Heitmann⁶², P. Hello²⁶, G. Hemming²⁸, M. Hendry⁴³, I. S. Heng⁴³, J. Hennig⁴³, J. Henry¹¹⁷, A. W. Heptonstall¹, M. Heurs^{9,21}, S. Hild⁴³, D. Hoak²⁸, D. Hofman²⁵, K. Holt⁶, D. E. Holz⁸⁶, P. Hopkins⁹⁶, C. Horst²⁰, J. Hough⁴³, E. A. Houston⁴³, E. J. Howell⁶⁰, Y. M. Hu⁹, E. A. Huerta¹¹, D. Huet²⁶, B. Hughey³³, S. Husa⁹⁷, S. H. Huttner⁴³, T. Huynh-Dinh⁶, N. Indik⁹, D. R. Ingram⁴⁴, R. Inta⁸⁰, G. Intini^{32,91}, H. N. Isa⁴³,

J.-M. Isac⁶⁸, M. Isi¹, B. R. Iyer¹⁹, K. Izumi⁴⁴, T. Jacqmin⁶⁸, K. Jani⁷³, P. Jaranowski¹¹⁸, S. Jawahar¹¹⁹, F. Jiménez-Forzeza⁹⁷, W. W. Johnson², D. I. Jones¹²⁰, R. Jones⁴³, R. J. G. Jonker¹³, L. Ju⁶⁰, J. Junker⁹, C. V. Kalaghatgi⁹⁶, V. Kalogera⁹⁵, S. Kandhasamy⁶, G. Kang³⁶, J. B. Kanner¹, S. Karki⁶⁷, K. S. Karvinen⁹, M. Kasprzak², M. Katolik¹¹, E. Katsavounidis¹⁴, W. Katzman⁶, S. Kaufer²¹, K. Kawabe⁴⁴, F. Kéfélian⁶², D. Keitel⁴³, A. J. Kembal¹¹, R. Kennedy¹⁰¹, C. Kent⁹⁶, J. S. Key¹²¹, F. Y. Khalili⁵⁷, I. Khan^{16,17}, S. Khan⁹, Z. Khan¹⁰⁰, E. A. Khazanov¹²², N. Kijbunchoo⁴⁴, Chunglee Kim¹²³, J. C. Kim¹²⁴, W. Kim⁷⁸, W. S. Kim¹²⁵, Y.-M. Kim^{87,123}, S. J. Kimbrell⁷³, E. J. King⁷⁸, P. J. King⁴⁴, R. Kirchhoff⁹, J. S. Kissel⁴⁴, L. Kleybolte³¹, S. Klimenko⁵, P. Koch⁹, S. M. Koehlenbeck⁹, S. Koley¹³, V. Kondrashov¹, A. Kontos¹⁴, M. Korobko³¹, W. Z. Korth¹, I. Kowalska⁷⁰, D. B. Kozak¹, C. Krämer⁹, V. Kringel⁹, B. Krishnan⁹, A. Królak^{126,127}, G. Kuehn⁹, P. Kumar⁸⁵, R. Kumar¹⁰⁰, S. Kumar¹⁹, L. Kuo⁸³, A. Kutynia¹²⁶, S. Kwang²⁰, B. D. Lackey³⁴, K. H. Lai⁸⁸, M. Landry⁴⁴, R. N. Lang²⁰, J. Lange¹¹⁷, B. Lantz⁴⁷, R. K. Lanza¹⁴, A. Lartaux-Vollard²⁶, P. D. Lasky¹⁰⁹, M. Laxen⁶, A. Lazzarini¹, C. Lazzaro⁵⁰, P. Leaci^{32,91}, S. Leavey⁴³, C. H. Lee⁸⁷, H. K. Lee¹²⁸, H. M. Lee¹²³, H. W. Lee¹²⁴, K. Lee⁴³, J. Lehmann⁹, A. Lenon^{37,38}, M. Leonardi^{90,104}, N. Leroy²⁶, N. Letendre⁷, Y. Levin¹⁰⁹, T. G. F. Li⁸⁸, A. Libson¹⁴, T. B. Littenberg¹²⁹, J. Liu⁶⁰, R. K. L. Lo⁸⁸, N. A. Lockerbie¹¹⁹, L. T. London⁹⁶, J. E. Lord⁴¹, M. Lorenzini^{16,17}, V. Lorette¹³⁰, M. Lormand⁶, G. Losurdo²³, J. D. Lough^{9,21}, C. O. Lousto¹¹⁷, G. Lovelace²⁷, H. Lück^{9,21}, D. Lumaca^{17,30}, A. P. Lundgren⁹, R. Lynch¹⁴, Y. Ma⁵⁹, S. Macfoy⁵⁸, B. Machenschalk⁹, M. MacInnis¹⁴, D. M. Macleod², I. Magaña Hernandez⁸⁸, F. Magaña-Sandoval⁴¹, L. Magaña Zertuche⁴¹, R. M. Magee⁸², E. Majorana³², I. Maksimovic¹³⁰, N. Man⁶², V. Mandic⁴², V. Mangano⁴³, G. L. Mansell²⁴, M. Manske²⁰, M. Mantovani²⁸, F. Marchesoni^{40,48}, F. Marion⁷, S. Márka⁴⁶, Z. Márka⁴⁶, C. Markakis¹¹, A. S. Markosyan⁴⁷, E. Maros¹, F. Martelli^{65,66}, L. Martellini⁶², I. W. Martin⁴³, D. V. Martynov¹⁴, K. Mason¹⁴, A. Masserot⁷, T. J. Massinger¹, M. Masso-Reid⁴³, S. Mastrogiovanni^{32,91}, A. Matas⁴², F. Matichard¹⁴, L. Matone⁴⁶, N. Mavalvala¹⁴, R. Mayani¹⁰², N. Mazumder⁶⁴, R. McCarthy⁴⁴, D. E. McClelland²⁴, S. McCormick⁶, L. McCuller¹⁴, S. C. McGuire¹³¹, G. McIntyre¹, J. McIver¹, D. J. McManus²⁴, T. McRae²⁴, S. T. McWilliams^{37,38}, D. Meacher⁸², G. D. Meadors^{9,34}, J. Meidam¹³, E. Mejuto-Villa⁸, A. Melatos¹³², G. Mendell⁴⁴, R. A. Mercer²⁰, E. L. Merilh⁴⁴, M. Merzougui⁶², S. Meshkov¹, C. Messenger⁴³, C. Messick⁸², R. Metzdorff⁶⁸, P. M. Meyers⁴², F. Mezzani^{32,91}, H. Miao⁵³, C. Michel²⁵, H. Middleton⁵³, E. E. Mikhailov¹³³, L. Milano^{4,75}, A. L. Miller⁵, A. Miller^{32,91}, B. B. Miller⁹⁵, J. Miller¹⁴, M. Millhouse⁹⁴, O. Minazzoli⁶², Y. Minenkov¹⁷, J. Ming³⁴, C. Mishra¹³⁴, S. Mitra¹⁸, V. P. Mitrofanov⁵⁷, G. Mitselmakher⁵, R. Mittleman¹⁴, A. Moggi²³, M. Mohan²⁸, S. R. P. Mohapatra¹⁴, M. Montani^{65,66}, B. C. Moore¹⁰⁵, C. J. Moore¹², D. Moraru⁴⁴, G. Moreno⁴⁴, S. R. Morriss⁹⁸, B. Mours⁷, C. M. Mow-Lowry⁵³, G. Mueller⁵, A. W. Muir⁹⁶, Arunava Mukherjee⁹, D. Mukherjee²⁰, S. Mukherjee⁹⁸, N. Mukund¹⁸, A. Mullavey⁶, J. Munch⁷⁸, E. A. M. Muniz⁴¹, P. G. Murray⁴³, K. Napier⁷³, I. Nardecchia^{17,30}, L. Naticchioni^{32,91}, R. K. Nayak¹³⁵, G. Nelemans^{13,61}, T. J. N. Nelson⁶, M. Neri^{54,55}, M. Nery⁹, A. Neunzert¹¹⁵, J. M. Newport¹¹⁶, G. Newton^{43,152}, K. K. Y. Ng⁸⁸, T. T. Nguyen²⁴, D. Nichols⁶¹, A. B. Nielsen⁹, S. Nisanke^{13,61}, A. Nitz⁹, A. Noack⁹, F. Nocera²⁸, D. Nolting⁶, M. E. N. Normandin⁹⁸, L. K. Nuttall⁴¹, J. Oberling⁴⁴, E. Ochsner²⁰, E. Oelker¹⁴, G. H. Ogil¹⁰⁷, J. J. Oh¹²⁵, S. H. Oh¹²⁵, F. Ohme⁹, M. Oliver⁹⁷, P. Oppermann⁹, Richard J. Oram⁶, B. O'Reilly⁶, R. Ormiston⁴², L. F. Ortega⁵, R. O'Shaughnessy¹¹⁷, D. J. Ottaway⁷⁸, H. Overmier⁶, B. J. Owen⁸⁰, A. E. Pace⁸², J. Page¹²⁹, M. A. Page⁶⁰, A. Pai¹¹¹, S. A. Pai⁵⁶, J. R. Palamos⁶⁷, O. Palashov¹²², C. Palomba³², A. Pal-Singh³¹, H. Pan⁸³, B. Pang⁵⁹, P. T. H. Pang⁸⁸, C. Pankow⁹⁵, F. Pannarale⁹⁶, B. C. Pant⁵⁶, F. Paoletti²³, A. Paoli²⁸, M. A. Papa^{9,20,34}, H. R. Paris⁴⁷, W. Parker⁶, D. Pascucci⁴³, A. Pasqualetti²⁸, R. Passaquietti^{22,23}, D. Passuello²³, B. Patricelli^{23,136}, B. L. Pearlstone⁴³, M. Pedraza¹, R. Pedurand^{25,137}, L. Pekowsky⁴¹, A. Pele⁶, S. Penn¹³⁸, C. J. Perez⁴⁴, A. Perreca^{1,90,104}, L. M. Perri⁹⁵, H. P. Pfeiffer⁸⁵, M. Phelps⁴³, O. J. Piccinni^{32,91}, M. Pichot⁶², F. Piergiovanni^{65,66}, V. Piero⁸, G. Pillant²⁸, L. Pinard²⁵, I. M. Pinto⁸, M. Pitkin⁴³, R. Poggiani^{22,23}, P. Popolizio²⁸, E. K. Porter³⁵, A. Post⁹, J. Powell⁴³, J. Prasad¹⁸, J. W. W. Pratt³³, V. Predoi⁹⁶, T. Prestegard²⁰, M. Prijatelj⁹, M. Principe⁸, S. Privitera³⁴, R. Prix⁹, G. A. Prodi^{90,104}, L. G. Prokhorov⁵⁷, O. Puncken⁹, M. Punturo⁴⁰, P. Puppò³², M. Pürer³⁴, H. Qi²⁰, J. Qin⁶⁰, S. Qiu¹⁰⁹, V. Quetschke⁹⁸, E. A. Quintero¹, R. Quitzow-James⁶⁷, F. J. Raab⁴⁴, D. S. Rabeling²⁴, H. Radkins⁴⁴, P. Raffai⁵¹, S. Raja⁵⁶, C. Rajan⁵⁶, M. Rakhmanov⁹⁸, K. E. Ramirez⁹⁸, P. Rapagnani^{32,91}, V. Raymond³⁴, M. Razzano^{22,23}, J. Read²⁷, T. Regimbau⁶², L. Rei⁵⁵, S. Reid⁵⁸, D. H. Reitze^{1,5}, H. Rew¹³³, S. D. Reyes⁴¹, F. Ricci^{32,91}, P. M. Ricker¹¹, S. Rieger⁹, K. Riles¹¹⁵, M. Rizzo¹¹⁷, N. A. Robertson^{1,43}, R. Robie⁴³, F. Robinet²⁶, A. Rocchi¹⁷, L. Rolland⁷, J. G. Rollins¹, V. J. Roma⁶⁷, R. Romano^{3,4}, C. L. Romel⁴⁴, J. H. Romie⁶, D. Rosińska^{52,139}, M. P. Ross¹⁴⁰, S. Rowan⁴³, A. Rüdiger⁹, P. Ruggi²⁸, K. Ryan⁴⁴, M. Rynge¹⁰², S. Sachdev¹, T. Sadecki⁴⁴, L. Sadeghian²⁰, M. Sakellariadou¹⁴¹, L. Salconi²⁸, M. Saleem¹¹¹, F. Salemi⁹, A. Samajdar¹³⁵, L. Sammut¹⁰⁹, L. M. Sampson⁹⁵, E. J. Sanchez¹, V. Sandberg⁴⁴, B. Sandeen⁹⁵, J. R. Sanders⁴¹, B. Sassolas²⁵, B. S. Sathyaprakash^{82,96}, P. R. Saulson⁴¹, O. Sauter¹¹⁵, R. L. Savage⁴⁴, A. Sawadsky²¹, P. Schale⁶⁷, J. Scheuer⁹⁵, E. Schmidt³³, J. Schmidt⁹, P. Schmidt^{1,61}, R. Schnabel³¹, R. M. S. Schofield⁶⁷, A. Schönbeck³¹, A. Schönbeck³¹, E. Schreiber⁹, D. Schuette^{9,21}, B. W. Schulte⁹, B. F. Schutz^{9,96}, S. G. Schwalbe³³, J. Scott⁴³, S. M. Scott²⁴, E. Seidel¹¹, D. Sellers⁶, A. S. Sengupta¹⁴², D. Sentenac²⁸, V. Sequino^{17,30}, A. Sergeev¹²², D. A. Shaddock²⁴, T. J. Shaffer⁴⁴, A. A. Shah¹²⁹, M. S. Shahriar⁹⁵, L. Shao³⁴, B. Shapiro⁴⁷, P. Shawhan⁷², A. Sheperd²⁰, D. H. Shoemaker¹⁴, D. M. Shoemaker⁷³, K. Siellez⁷³, X. Siemens²⁰, M. Sieniawska⁵², D. Sigg⁴⁴, A. D. Silva¹⁵, A. Singer¹, L. P. Singer⁷⁶, A. Singh^{9,21,34}, R. Singh², A. Singhal^{16,32}, A. M. Sintes⁹⁷, B. J. J. Slagmolen²⁴, B. Smith⁶, J. R. Smith²⁷, R. J. E. Smith¹, E. J. Son¹²⁵, J. A. Sonnenberg²⁰, B. Sorazu⁴³, F. Sorrentino⁵⁵, T. Souradeep¹⁸, A. P. Spencer⁴³, A. K. Srivastava¹⁰⁰, A. Staley⁴⁶, M. Steinke⁹, J. Steinlechner^{31,43}, S. Steinlechner³¹, D. Steinmeyer^{9,21}, B. C. Stephens²⁰, R. Stone⁹⁸, K. A. Strain⁴³, G. Stratta^{65,66}, S. E. Strigin⁵⁷, R. Sturani¹⁴³, A. L. Stuver⁶, T. Z. Summerscales¹⁴⁴, L. Sun¹³², S. Sunil¹⁰⁰, P. J. Sutton⁹⁶, B. L. Swinkels²⁸, M. J. Szczepańczyk³³, M. Tacca³⁵, D. Talukder⁶⁷, D. B. Tanner⁵, M. Tápai¹¹², A. Taracchini³⁴, J. A. Taylor¹²⁹, R. Taylor¹, T. Theeg⁹, E. G. Thomas⁵³, M. Thomas⁶, P. Thomas⁴⁴, K. A. Thorne⁶, K. S. Thorne⁵⁹, E. Thrane¹⁰⁹

S. Tiwari^{16,90}, V. Tiwari⁹⁶, K. V. Tokmakov¹¹⁹, K. Toland⁴³, M. Tonelli^{22,23}, Z. Tornasi⁴³, C. I. Torrie¹, D. Töyrä⁵³, F. Traverso^{28,40}, G. Traylor⁶, D. Trifirò¹⁰, J. Trinastic⁵, M. C. Tringali^{90,104}, L. Trozzo^{23,145}, K. W. Tsang¹³, M. Tse¹⁴, R. Tso¹, D. Tuyenbayev⁹⁸, K. Ueno²⁰, D. Ugolini¹⁴⁶, C. S. Unnikrishnan¹¹³, A. L. Urban¹, S. A. Usman⁹⁶, K. Vahi¹⁰², H. Vahlbruch²¹, G. Vajente¹, G. Valdes⁹⁸, M. Vallisneri⁵⁹, N. van Bakel¹³, M. van Beuzekom¹³, J. F. J. van den Brand^{13,71}, C. Van Den Broeck¹³, D. C. Vander-Hyde⁴¹, L. van der Schaaf¹³, J. V. van Heijningen¹³, A. A. van Veggel⁴³, M. Vardaro^{49,50}, V. Varma⁵⁹, S. Vass¹, M. Vasúth⁴⁵, A. Vecchio⁵³, G. Vedovato⁵⁰, J. Veitch⁵³, P. J. Veitch⁷⁸, K. Venkateswara¹⁴⁰, G. Venugopalan¹, D. Verkindt⁷, F. Vetranò^{65,66}, A. Vicere^{65,66}, A. D. Viets²⁰, S. Vinciguerra⁵³, D. J. Vine⁵⁸, J.-Y. Vinet⁶², S. Vitale¹⁴, T. Vo⁴¹, H. Vocca^{39,40}, C. Vorvick⁴⁴, D. V. Voss⁵, W. D. Vousden⁵³, S. P. Vyatchanin⁵⁷, A. R. Wade¹, L. E. Wade⁸¹, M. Wade⁸¹, R. Walet¹³, M. Walker², L. Wallace¹, S. Walsh²⁰, G. Wang^{16,66}, H. Wang⁵³, J. Z. Wang⁸², M. Wang⁵³, Y.-F. Wang⁸⁸, Y. Wang⁶⁰, R. L. Ward²⁴, J. Warner⁴⁴, M. Was⁷, J. Watchi⁹², B. Weaver⁴⁴, L.-W. Wei^{9,21}, M. Weinert⁹, A. J. Weinstein¹, R. Weiss¹⁴, L. Wen⁶⁰, E. K. Wessel¹¹, P. Wessels⁹, T. Westphal⁹, K. Wette⁹, J. T. Whelan¹¹⁷, B. F. Whiting⁵, C. Whittle¹⁰⁹, D. Williams⁴³, R. D. Williams¹, A. R. Williamson¹¹⁷, J. L. Willis¹⁴⁷, B. Willke^{9,21}, M. H. Wimmer^{9,21}, W. Winkler⁹, C. C. Wipf¹, H. Wittel^{9,21}, G. Woan⁴³, J. Woehler⁹, J. Wofford¹¹⁷, K. W. K. Wong⁸⁸, J. Worden⁴⁴, J. L. Wright⁴³, D. S. Wu⁹, G. Wu⁶, W. Yam¹⁴, H. Yamamoto¹, C. C. Yancey⁷², M. J. Yap²⁴, Hang Yu¹⁴, Haocun Yu¹⁴, M. Yvert⁷, M. Zanolin³³, T. Zelenova²⁸, J.-P. Zendri⁵⁰, M. Zevin⁹⁵, L. Zhang¹, M. Zhang¹³³, T. Zhang⁴³, Y.-H. Zhang¹¹⁷, C. Zhao⁶⁰, M. Zhou⁹⁵, Z. Zhou⁹⁵, X. J. Zhu⁶⁰, M. E. Zucker^{1,14}, J. Zweizig¹

LIGO Scientific Collaboration and Virgo Collaboration,

D. Steeghs^{109,148}, and L. Wang¹⁴⁸

¹ LIGO, California Institute of Technology, Pasadena, CA 91125, USA

² Louisiana State University, Baton Rouge, LA 70803, USA

³ Università di Salerno, Fisciano, I-84084 Salerno, Italy

⁴ INFN, Sezione di Napoli, Complesso Universitario di Monte S. Angelo, I-80126 Napoli, Italy

⁵ University of Florida, Gainesville, FL 32611, USA

⁶ LIGO Livingston Observatory, Livingston, LA 70754, USA

⁷ Laboratoire d'Annecy-le-Vieux de Physique des Particules (LAPP), Université Savoie Mont Blanc, CNRS/IN2P3, F-74941 Annecy, France

⁸ University of Sannio at Benevento, I-82100 Benevento, Italy and INFN, Sezione di Napoli, I-80100 Napoli, Italy

⁹ Albert-Einstein-Institut, Max-Planck-Institut für Gravitationsphysik, D-30167 Hannover, Germany

¹⁰ The University of Mississippi, University, MS 38677, USA

¹¹ NCSA, University of Illinois at Urbana-Champaign, Urbana, IL 61801, USA

¹² University of Cambridge, Cambridge CB2 1TN, UK

¹³ Nikhef, Science Park, 1098 XG Amsterdam, The Netherlands

¹⁴ LIGO, Massachusetts Institute of Technology, Cambridge, MA 02139, USA

¹⁵ Instituto Nacional de Pesquisas Espaciais, 12227-010 São José dos Campos, São Paulo, Brazil

¹⁶ Gran Sasso Science Institute (GSSI), I-67100 L'Aquila, Italy

¹⁷ INFN, Sezione di Roma Tor Vergata, I-00133 Roma, Italy

¹⁸ Inter-University Centre for Astronomy and Astrophysics, Pune 411007, India

¹⁹ International Centre for Theoretical Sciences, Tata Institute of Fundamental Research, Bengaluru 560089, India

²⁰ University of Wisconsin-Milwaukee, Milwaukee, WI 53201, USA

²¹ Leibniz Universität Hannover, D-30167 Hannover, Germany

²² Università di Pisa, I-56127 Pisa, Italy

²³ INFN, Sezione di Pisa, I-56127 Pisa, Italy

²⁴ OzGrav, Australian National University, Canberra, Australian Capital Territory 0200, Australia

²⁵ Laboratoire des Matériaux Avancés (LMA), CNRS/IN2P3, F-69622 Villeurbanne, France

²⁶ LAL, Univ. Paris-Sud, CNRS/IN2P3, Université Paris-Saclay, F-91898 Orsay, France

²⁷ California State University Fullerton, Fullerton, CA 92831, USA

²⁸ European Gravitational Observatory (EGO), I-56021 Cascina, Pisa, Italy

²⁹ Chennai Mathematical Institute, Chennai 603103, India

³⁰ Università di Roma Tor Vergata, I-00133 Roma, Italy

³¹ Universität Hamburg, D-22761 Hamburg, Germany

³² INFN, Sezione di Roma, I-00185 Roma, Italy

³³ Embry-Riddle Aeronautical University, Prescott, AZ 86301, USA

³⁴ Albert-Einstein-Institut, Max-Planck-Institut für Gravitationsphysik, D-14476 Potsdam-Golm, Germany

³⁵ APC, AstroParticule et Cosmologie, Université Paris Diderot, CNRS/IN2P3, CEA/Irfu, Observatoire de Paris, Sorbonne Paris Cité, F-75205 Paris Cedex 13, France

³⁶ Korea Institute of Science and Technology Information, Daejeon 34141, Korea

³⁷ West Virginia University, Morgantown, WV 26506, USA

³⁸ Center for Gravitational Waves and Cosmology, West Virginia University, Morgantown, WV 26505, USA

³⁹ Università di Perugia, I-06123 Perugia, Italy

⁴⁰ INFN, Sezione di Perugia, I-06123 Perugia, Italy

⁴¹ Syracuse University, Syracuse, NY 13244, USA

⁴² University of Minnesota, Minneapolis, MN 55455, USA

⁴³ SUPA, University of Glasgow, Glasgow G12 8QQ, UK

⁴⁴ LIGO Hanford Observatory, Richland, WA 99352, USA

⁴⁵ Wigner RCP, RMKI, H-1121 Budapest, Konkoly Thege Miklós út 29-33, Hungary

⁴⁶ Columbia University, New York, NY 10027, USA

⁴⁷ Stanford University, Stanford, CA 94305, USA

⁴⁸ Università di Camerino, Dipartimento di Fisica, I-62032 Camerino, Italy

⁴⁹ Università di Padova, Dipartimento di Fisica e Astronomia, I-35131 Padova, Italy

⁵⁰ INFN, Sezione di Padova, I-35131 Padova, Italy

⁵¹ MTA Eötvös University, "Lendulet" Astrophysics Research Group, Budapest 1117, Hungary

⁵² Nicolaus Copernicus Astronomical Center, Polish Academy of Sciences, 00-716, Warsaw, Poland

⁵³ University of Birmingham, Birmingham B15 2TT, UK

- ⁵⁴ Università degli Studi di Genova, I-16146 Genova, Italy
⁵⁵ INFN, Sezione di Genova, I-16146 Genova, Italy
⁵⁶ RRCAT, Indore MP 452013, India
⁵⁷ Faculty of Physics, Lomonosov Moscow State University, Moscow 119991, Russia
⁵⁸ SUPA, University of the West of Scotland, Paisley PA1 2BE, UK
⁵⁹ Caltech CaRT, Pasadena, CA 91125, USA
⁶⁰ OzGrav, University of Western Australia, Crawley, Western Australia 6009, Australia
⁶¹ Department of Astrophysics/IMAPP, Radboud University Nijmegen, P.O. Box 9010, 6500 GL Nijmegen, The Netherlands
⁶² Artemis, Université Côte d'Azur, Observatoire Côte d'Azur, CNRS, CS 34229, F-06304 Nice Cedex 4, France
⁶³ Institut de Physique de Rennes, CNRS, Université de Rennes 1, F-35042 Rennes, France
⁶⁴ Washington State University, Pullman, WA 99164, USA
⁶⁵ Università degli Studi di Urbino "Carlo Bo," I-61029 Urbino, Italy
⁶⁶ INFN, Sezione di Firenze, I-50019 Sesto Fiorentino, Firenze, Italy
⁶⁷ University of Oregon, Eugene, OR 97403, USA
⁶⁸ Laboratoire Kastler Brossel, UPMC-Sorbonne Universités, CNRS, ENS-PSL Research University, Collège de France, F-75005 Paris, France
⁶⁹ Carleton College, Northfield, MN 55057, USA
⁷⁰ Astronomical Observatory Warsaw University, 00-478 Warsaw, Poland
⁷¹ VU University Amsterdam, 1081 HV Amsterdam, The Netherlands
⁷² University of Maryland, College Park, MD 20742, USA
⁷³ Center for Relativistic Astrophysics and School of Physics, Georgia Institute of Technology, Atlanta, GA 30332, USA
⁷⁴ Université Claude Bernard Lyon 1, F-69622 Villeurbanne, France
⁷⁵ Università di Napoli "Federico II", Complesso Universitario di Monte S. Angelo, I-80126 Napoli, Italy
⁷⁶ NASA Goddard Space Flight Center, Greenbelt, MD 20771, USA
⁷⁷ RESCEU, University of Tokyo, Tokyo, 113-0033, Japan
⁷⁸ OzGrav, University of Adelaide, Adelaide, South Australia 5005, Australia
⁷⁹ Tsinghua University, Beijing 100084, China
⁸⁰ Texas Tech University, Lubbock, TX 79409, USA
⁸¹ Kenyon College, Gambier, OH 43022, USA
⁸² The Pennsylvania State University, University Park, PA 16802, USA
⁸³ National Tsing Hua University, Hsinchu City, 30013 Taiwan, Republic of China
⁸⁴ Charles Sturt University, Wagga Wagga, New South Wales 2678, Australia
⁸⁵ Canadian Institute for Theoretical Astrophysics, University of Toronto, Toronto, Ontario M5S 3H8, Canada
⁸⁶ University of Chicago, Chicago, IL 60637, USA
⁸⁷ Pusan National University, Busan 46241, Korea
⁸⁸ The Chinese University of Hong Kong, Shatin, NT, Hong Kong
⁸⁹ INAF, Osservatorio Astronomico di Padova, Vicolo dell'Osservatorio 5, I-35122 Padova, Italy
⁹⁰ INFN, Trento Institute for Fundamental Physics and Applications, I-38123 Povo, Trento, Italy
⁹¹ Università di Roma "La Sapienza", I-00185 Roma, Italy
⁹² Université Libre de Bruxelles, Brussels 1050, Belgium
⁹³ Sonoma State University, Rohnert Park, CA 94928, USA
⁹⁴ Montana State University, Bozeman, MT 59717, USA
⁹⁵ Center for Interdisciplinary Exploration & Research in Astrophysics (CIERA), Northwestern University, Evanston, IL 60208, USA
⁹⁶ Cardiff University, Cardiff CF24 3AA, UK
⁹⁷ Universitat de les Illes Balears, IAC3—IEEC, E-07122 Palma de Mallorca, Spain
⁹⁸ The University of Texas Rio Grande Valley, Brownsville, TX 78520, USA
⁹⁹ Bellevue College, Bellevue, WA 98007, USA
¹⁰⁰ Institute for Plasma Research, Bhat, Gandhinagar 382428, India
¹⁰¹ The University of Sheffield, Sheffield S10 2TN, UK
¹⁰² University of Southern California Information Sciences Institute, Marina Del Rey, CA 90292, USA
¹⁰³ California State University, Los Angeles, 5151 State University Dr, Los Angeles, CA 90032, USA
¹⁰⁴ Università di Trento, Dipartimento di Fisica, I-38123 Povo, Trento, Italy
¹⁰⁵ Montclair State University, Montclair, NJ 07043, USA
¹⁰⁶ National Astronomical Observatory of Japan, 2-21-1 Osawa, Mitaka, Tokyo 181-8588, Japan
¹⁰⁷ Whitman College, 345 Boyer Avenue, Walla Walla, WA 99362 USA
¹⁰⁸ School of Mathematics, University of Edinburgh, Edinburgh EH9 3FD, UK
¹⁰⁹ OzGrav, School of Physics & Astronomy, Monash University, Clayton 3800, Victoria, Australia
¹¹⁰ University and Institute of Advanced Research, Gandhinagar Gujarat 382007, India
¹¹¹ IISER-TVM, CET Campus, Trivandrum Kerala 695016, India
¹¹² University of Szeged, Dóm tér 9, Szeged 6720, Hungary
¹¹³ Tata Institute of Fundamental Research, Mumbai 400005, India
¹¹⁴ INAF, Osservatorio Astronomico di Capodimonte, I-80131, Napoli, Italy
¹¹⁵ University of Michigan, Ann Arbor, MI 48109, USA
¹¹⁶ American University, Washington, DC 20016, USA
¹¹⁷ Rochester Institute of Technology, Rochester, NY 14623, USA
¹¹⁸ University of Bia lystok, 15-424 Bia lystok, Poland
¹¹⁹ SUPA, University of Strathclyde, Glasgow G1 1XQ, UK
¹²⁰ University of Southampton, Southampton SO17 1BJ, UK
¹²¹ University of Washington Bothell, 18115 Campus Way NE, Bothell, WA 98011, USA
¹²² Institute of Applied Physics, Nizhny Novgorod, 603950, Russia
¹²³ Seoul National University, Seoul 08826, Korea
¹²⁴ Inje University Gimhae, South Gyeongsang 50834, Korea
¹²⁵ National Institute for Mathematical Sciences, Daejeon 34047, Korea
¹²⁶ NCBJ, 05-400 Świerk-Otwock, Poland
¹²⁷ Institute of Mathematics, Polish Academy of Sciences, 00656 Warsaw, Poland
¹²⁸ Hanyang University, Seoul 04763, Korea
¹²⁹ NASA Marshall Space Flight Center, Huntsville, AL 35811, USA

- ¹³⁰ ESPCI, CNRS, F-75005 Paris, France
¹³¹ Southern University and A&M College, Baton Rouge, LA 70813, USA
¹³² OzGrav, University of Melbourne, Parkville, Victoria 3010, Australia
¹³³ College of William and Mary, Williamsburg, VA 23187, USA
¹³⁴ Indian Institute of Technology Madras, Chennai 600036, India
¹³⁵ IISER-Kolkata, Mohanpur, West Bengal 741252, India
¹³⁶ Scuola Normale Superiore, Piazza dei Cavalieri 7, I-56126 Pisa, Italy
¹³⁷ Université de Lyon, F-69361 Lyon, France
¹³⁸ Hobart and William Smith Colleges, Geneva, NY 14456, USA
¹³⁹ Janusz Gil Institute of Astronomy, University of Zielona Góra, 65-265 Zielona Góra, Poland
¹⁴⁰ University of Washington, Seattle, WA 98195, USA
¹⁴¹ King's College London, University of London, London WC2R 2LS, UK
¹⁴² Indian Institute of Technology, Gandhinagar Ahmedabad Gujarat 382424, India
¹⁴³ International Institute of Physics, Universidade Federal do Rio Grande do Norte, Natal RN 59078-970, Brazil
¹⁴⁴ Andrews University, Berrien Springs, MI 49104, USA
¹⁴⁵ Università di Siena, I-53100 Siena, Italy
¹⁴⁶ Trinity University, San Antonio, TX 78212, USA
¹⁴⁷ Abilene Christian University, Abilene, TX 79699, USA
¹⁴⁸ Department of Physics, University of Warwick, Gibbet Hill Road, Coventry CV4 7AL, UK
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Abstract

We present the results of a semicoherent search for continuous gravitational waves from the low-mass X-ray binary Scorpius X-1, using data from the first Advanced LIGO observing run. The search method uses details of the modeled, parametrized continuous signal to combine coherently data separated by less than a specified coherence time, which can be adjusted to trade off sensitivity against computational cost. A search was conducted over the frequency range 25–2000 Hz, spanning the current observationally constrained range of binary orbital parameters. No significant detection candidates were found, and frequency-dependent upper limits were set using a combination of sensitivity estimates and simulated signal injections. The most stringent upper limit was set at 175 Hz, with comparable limits set across the most sensitive frequency range from 100 to 200 Hz. At this frequency, the 95% upper limit on the signal amplitude h_0 is 2.3×10^{-25} marginalized over the unknown inclination angle of the neutron star's spin, and 8.0×10^{-26} assuming the best orientation (which results in circularly polarized gravitational waves). These limits are a factor of 3–4 stronger than those set by other analyses of the same data, and a factor of ~ 7 stronger than the best upper limits set using data from Initial LIGO science runs. In the vicinity of 100 Hz, the limits are a factor of between 1.2 and 3.5 above the predictions of the torque balance model, depending on the inclination angle; if the most likely inclination angle of 44° is assumed, they are within a factor of 1.7.

Key words: accretion, accretion disks – gravitational waves – stars: neutron – X-rays: binaries

1. Introduction

Rotating neutron stars (NSs) are the primary expected source of continuous, periodic gravitational waves (GWs) for ground-based GW detectors. Targets include known pulsars (Aasi et al. 2014a), non-pulsating NSs in supernova remnants (Wette et al. 2008; Abadie et al. 2010; Aasi et al. 2015a), and unknown isolated (Aasi et al. 2016; Abbott et al. 2016a) or binary NSs (Aasi et al. 2014b). A particularly promising source is an accreting NS in a low-mass X-ray binary (LMXB); accretion torque spins up the NS into the frequency band of the detectors, and the accretion can generate an asymmetric mass or current quadrupole that acts as the source for the GWs (Watts et al. 2008). An approximate equilibrium between the accretion spin-up and GW spin-down, as well as other spin-down torques can produce a signal that is nearly periodic in the NS's rest frame,

and then Doppler-shifted due to the orbital motion of the NS and the motion of the detector on the surface of the Earth. Such an equilibrium scenario would produce a relation between the observed accretion-induced X-ray flux of the LMXB and the expected strength of the GWs. Scorpius X-1 (Sco X-1), the most luminous LMXB, is therefore a promising potential source of GWs (Papaloizou & Pringle 1978; Wagoner 1984; Bildsten 1998). Sco X-1 is presumed to consist of an NS of mass $\approx 1.4 M_\odot$ in a binary orbit with a companion star of mass $\approx 0.4 M_\odot$ (Steeghs & Casares 2002). Some of the parameters inferred from observations of the system are summarized in Table 1.

Several methods were used to search for Sco X-1 in data from the Initial LIGO science runs of 2002–2011: Abbott et al. (2007a) performed a fully coherent search (Jaranowski et al. 1998) on six hours of data from the second science run. Starting with the fourth science run, results for Sco X-1 were reported (Abbott et al. 2007b; Abadie et al. 2011) as part of a search for stochastic signals from isolated sky positions (Ballmer 2006). In the fifth science run, a search (Aasi et al. 2015b) was done for Doppler-modulated sidebands associated with the binary orbit (Messenger & Woan 2007; Sammut et al. 2014). In the sixth science run, Sco X-1 was included in a search (Aasi et al. 2014b) principally designed for unknown binary systems (Goetz & Riles 2011), and this method

¹⁴⁹ Deceased, March 2016.

¹⁵⁰ Deceased, March 2017.

¹⁵¹ Deceased, February 2017.

¹⁵² Deceased, December 2016.



Table 1
Observed Parameters of the LMXB Sco X-1

| Parameter | Value |
|--|---|
| R.A. ^a | 16 ^h 19 ^m 55 ^s .0850 |
| Decl. ^a | −15°38′24″.9 |
| Distance (kpc) | 2.8 ± 0.3 |
| Orbital inclination i^b | 44° ± 6° |
| K_1 (km s ^{−1}) ^c | [10, 90] or [40, 90] |
| T_{asc} (GPS s) ^d | 897753994 ± 100 |
| P_{orb} (s) ^d | 68023.70 ± 0.04 |

Notes. Uncertainties are 1σ unless otherwise stated. There are uncertainties (relevant to the present search) in the projected velocity amplitude K_1 of the NS, the orbital period P_{orb} , and the time T_{asc} at which the neutron star crosses the ascending node (moving away from the observer), measured in the solar system barycenter. The orbital eccentricity of Sco X-1 is believed to be small (Steehgs & Casares 2002; Wang 2017) and is ignored in this search. The inclusion of eccentric orbits would add two search parameters that are determined by the eccentricity and the argument of periaipse (Messenger 2011; Leaci & Prix 2015).

^a The sky position (as quoted in Abbott et al. 2007a and derived from Bradshaw et al. 1999) is determined to the microarcsecond, and therefore can be treated as known in the present search.

^b The inclination i of the orbit to the line of sight, from the observation of radio jets in Fomalont et al. (2001), is not necessarily the same as the inclination angle ι of the neutron star’s spin axis, which determines the degree of polarization of the GW in Equation (1).

^c The value of the projected orbital velocity K_1 is difficult to determine experimentally, and previous works used a value from Abbott et al. (2007a), which was derived with some assumptions from Steehgs & Casares (2002) and equivalent to 40 ± 5 km s^{−1}. The broader range listed here comes from Doppler tomography measurements and Monte Carlo simulations in Wang (2017), which show K_1 to be weakly determined beyond the constraint that $40 \text{ km s}^{-1} \lesssim K_1 \lesssim 90 \text{ km s}^{-1}$. Preliminary results from Wang (2017) included the weaker constraint $10 \text{ km s}^{-1} \lesssim K_1 \lesssim 90 \text{ km s}^{-1}$, which was used to determine the parameter range in Table 2.

^d The time of ascension T_{asc} , at which the neutron star crosses the ascending node (moving away from the observer), measured in the solar system barycenter, is derived from the time of inferior conjunction of the companion given in Galloway et al. (2014) by subtracting $P_{\text{orb}}/4$. It corresponds to a time of 2008 June 17 16:06:20 UTC and can be propagated to other epochs by adding an integer multiple of P_{orb} , which results in increased uncertainty in T_{asc} and correlations between P_{orb} and T_{asc} ; see Figure 1.

References. Bradshaw et al. (1999), Fomalont et al. (2001), Galloway et al. (2014), Wang (2017).

was subsequently improved to search directly for Sco X-1 (Meadors et al. 2016) and applied to Initial LIGO data (Meadors et al. 2017). A mock data challenge (Messenger et al. 2015) was conducted to compare several of the methods to search for Sco X-1, and the most sensitive (detecting all 50 simulated signals in the challenge, and 49 out of the 50 “training” signals) was the cross-correlation (CrossCorr) method (Dhurandhar et al. 2008; Whelan et al. 2015) used in the present analysis.¹⁵³

The Advanced LIGO detectors (Aasi et al. 2015c) carried out their first observing run (O1) from 2015 September 12 to 2016 January 19 (Abbott et al. 2016b). Searches for transient signals were carried out in near-real time and resulted in the observation of the binary black hole (BBH) mergers GW 150914 (Abbott et al. 2016c) and GW 151226 (Abbott et al. 2016d), and the

¹⁵³ The CrossCorr analysis was carried out in “self-blinded” mode without knowledge of the simulated signal parameters, after the nominal end of the challenge.

possible BBH merger LVT 151012 (Abbott et al. 2016b), as well as upper limits on the rates and strengths of other sources (Abbott et al. 2016e, 2017a, 2017b). Searches for persistent stochastic or periodic sources were conducted using data from the full duration of the run and include searches for isotropic and anisotropic stochastic signals (Abbott et al. 2017c, 2017d) and a variety of known and unknown NSs (Abbott et al. 2017e). So far, two analyses including searches for GWs from Sco X-1 besides the current one have been released: Abbott et al. (2017d) included the direction of Sco X-1 in their directed unmodeled search for persistent GWs, and Abbott et al. (2017f) performed a directed search for Sco X-1 using a hidden Markov model.

2. Model of GWs from Sco X-1

The modeled GW signal from a rotating NS consists of a “plus” polarization component, $h_+(t) = A_+ \cos[\Phi(t)]$, and a “cross” polarization component, $h_\times(t) = A_\times \sin[\Phi(t)]$. The signal recorded in a particular detector will be a linear combination of h_+ and h_\times determined by the detector’s orientation as a function of time. The two polarization amplitudes are

$$A_+ = h_0 \frac{1 + \cos^2 \iota}{2} \quad \text{and} \quad A_\times = h_0 \cos \iota, \quad (1)$$

where h_0 is an intrinsic amplitude related to the NS’s ellipticity, moment of inertia, spin frequency, and distance; and ι is the inclination of the NS’s spin to the line of sight. (For an NS in a binary, this may or may not be related to the inclination i of the binary orbit.) If $\iota = 0^\circ$ or 180° , $A_\times = \pm A_+$, and gravitational radiation is circularly polarized. If $\iota = 90^\circ$, $A_\times = 0$, it is linearly polarized. The general case, elliptical polarization, has $0 < |A_\times| < A_+$. Many search methods are sensitive to the combination

$$(h_0^{\text{eff}})^2 = \frac{A_+^2 + A_\times^2}{2} = h_0^2 \frac{[(1 + \cos^2 \iota)/2]^2 + [\cos \iota]^2}{2}, \quad (2)$$

which is equal to h_0^2 for circular polarization and $h_0^2/8$ for linear polarization (Messenger et al. 2015; note that this differs by a factor of 2.5 from the definition of $(h_0^{\text{eff}})^2$ in Whelan et al. 2015).

It has been suggested (Papaloizou & Pringle 1978; Wagoner 1984; Bildsten 1998) that an LMXB may be in an equilibrium state where the spin-up due to accretion is due to the spin-down due to GWs. In that case, the GW amplitude can be related to the accretion rate, as inferred from the X-ray flux F_X (Watts et al. 2008):

$$h_0 \approx 3 \times 10^{-27} \left(\frac{F_X}{10^{-8} \text{ erg cm}^{-2} \text{ s}^{-1}} \right)^{1/2} \left(\frac{\nu_s}{300 \text{ Hz}} \right)^{-1/2} \times \left(\frac{R}{10 \text{ km}} \right)^{3/4} \left(\frac{M}{1.4 M_\odot} \right)^{-1/4}.$$

For Sco X-1, using the observed X-ray flux $F_X = 3.9 \times 10^{-7} \text{ erg cm}^{-2} \text{ s}^{-1}$ from Watts et al. (2008), and assuming that the GW frequency f_0 is twice the spin frequency ν_s (as would be the case for GWs generated by triaxiality in the NS), the torque

balance value is

$$h_0 \approx 3.4 \times 10^{-26} \left(\frac{f_0}{600 \text{ Hz}} \right)^{-1/2}. \quad (4)$$

Recent works (Haskell et al. 2015a, 2015b) have cast doubt on the ubiquity of the GW torque balance scenario in light of other spin-down mechanisms; the torque balance level remains an important benchmark for search sensitivity, and the detection or non-detection at or below that level would provide insight into the behavior of accreting NSs.

3. CrossCorr Search Method

The CrossCorr method was presented in Dhurandhar et al. (2008) and refined for application to Sco X-1 in Whelan et al. (2015). It was applied to simulated Advanced LIGO data in a mock data challenge (Messenger et al. 2015; Y. Zhang et al. 2017, in preparation). It was originally developed as a model-based improvement of the directional stochastic search of Ballmer (2006), which has been used to set limits on gravitational radiation from specific sky directions including Sco X-1 (Abbott et al. 2007b; Abadie et al. 2011). The method allows data to be correlated up to an adjustable coherence time T_{max} . The data are split into segments of length T_{fit} between 240 and 1400 s (depending on frequency) and Fourier transformed. In a given data segment or short Fourier transform (SFT), the signal is expected to be found in a particular Fourier bin (or bins, considering the effects of spectral leakage). The signal bins are determined by the intrinsic frequency and the expected Doppler shift, which is in turn determined by the time and detector location, as well as the assumed orbital parameters of the LMXB. If the SFTs are labelled by the index K , L , etc., which encodes both the detector in question and the time of the SFT, and z_K is the appropriately normalized Fourier data in the bin(s) of interest, the CrossCorr statistic has the form

$$\rho = \sum_{KL \in \mathcal{P}} (W_{KL} z_K^* z_L + W_{KL}^* z_K z_L^*). \quad (5)$$

This includes the product of the data from SFTs K and L , where KL is in a list of allowed pairs \mathcal{P} , defined by $K < L$ and $|T_K - T_L| \leq T_{\text{max}}$, i.e., the times of the two different data segments should differ by no more than some specified lag time T_{max} , which we also refer to as the coherence time. The complex weighting factors W_{KL} are chosen (according to Equations (2.33)–(2.36) and (3.5) of Whelan et al. 2015) to maximize the expected statistic value subject to the normalization $\text{Var}(\rho) = 1$. The expected statistic value is then

$$E[\rho] = (h_0^{\text{eff}})^2 \vartheta, \quad (6)$$

where

$$\vartheta \approx 0.903 \sqrt{N_{\text{det}}^2 T_{\text{obs}} T_{\text{max}} \left\langle \frac{4(\Gamma_{KL}^{\text{ave}})^2}{S_K S_L} \right\rangle_{KL \in \mathcal{P}}}, \quad (7)$$

(this is the quantity called ϱ^{ave} in Whelan et al. 2015) and h_0^{eff} is the combination of h_0 and $\cos i$ defined in Equation (2), S_K is constructed from the noise power spectrum and Γ_{KL}^{ave} from the antenna patterns for detectors K and L at the appropriate times, N_{det} is the number of detectors participating in the search, T_{obs} is the observing time per detector, and the factor of 0.903 arises from spectral leakage, assuming we consider contributions

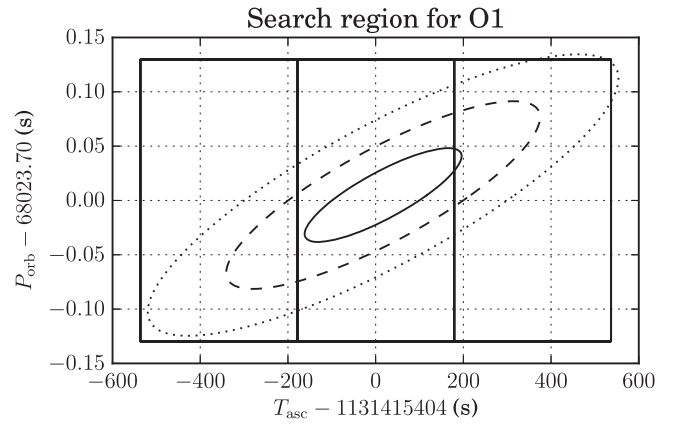


Figure 1. Range of search parameters T_{asc} and P_{orb} . The ellipses show curves of constant prior probability corresponding to 1σ , 2σ , and 3σ (containing 39.3%, 86.5%, and 98.9% of the prior probability, respectively), including the effect of correlations arising from the propagation of the T_{asc} estimate from Table 1 to the mid-run value in Table 2. The search region is chosen to include the 3σ ellipse, with the range of T_{asc} within $\pm 1\sigma$ receiving a deeper search, as illustrated in Figure 2. The inner and outer regions contain 68.0% and 99.5% of the prior probability, respectively. Note that the apparent inefficiency in searching unlikely regions of $T_{\text{asc}}-P_{\text{orb}}$ space is mitigated by the fact that the search does not typically resolve P_{orb} , resulting in only one value being included in the search.

from all Fourier bins. (See Equation (3.19) of Whelan et al. 2015 for more details.) Increasing T_{max} increases the sensitivity of the search, but also increases the computing cost. In order to maximize the chance for a potential detection, a range of choices for T_{max} was used for different values of signal frequency and orbital parameters. The method used longer coherence times in regions of parameter space where (1) the detectable signal level given the frequency-dependent instrumental noise was closer to the expected signal strength from torque balance, (2) the cost of the search was lower due to template spacing, i.e., at lower frequencies and $a \sin i$ values, or (3) the signal had higher prior probability of being found, i.e., closer to the most likely value of T_{asc} . This is illustrated in Figure 2. The full set of coherence times used ranges from 25,290 s for 25–50 Hz (for the most likely T_{asc} and smallest $a \sin i$ values) to 240 s at frequencies above 1200 Hz.

The search was performed using a bank of template signals laid out in hypercubic lattice in the signal parameters of intrinsic frequency f_0 , projected semimajor axis $a \sin i$, time of ascension T_{asc} , and (where appropriate) orbital period P_{orb} . The range of values in each direction, motivated by Table 1 and Figure 1, is shown in Table 2. The lattice spacing for the initial search was chosen to correspond to a nominal metric mismatch (fractional loss of signal-to-noise ratio (S/N) associated with a one-lattice-spacing offset in a given direction, assuming quadratic approximation) of 25% in each of the four parameters, using the metric computed in Whelan et al. (2015). The lattice was constructed (and spacing computed) for each of the 18 orbital parameter space cells shown in Figure 2 in each 0.05 Hz-wide frequency band. This resulted in a total of $\sim 9 \times 10^5 - 2 \times 10^8$ detection statistics per 0.05 Hz, as detailed in Table 3.

4. Follow-up of Candidates

Although the detection statistic ρ is normalized to have zero mean and unit variance in Gaussian noise, the trials factor

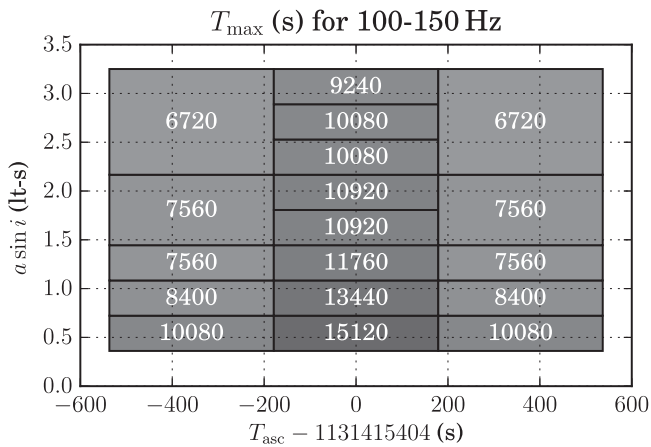


Figure 2. Example of coherence times T_{\max} , in seconds, chosen as a function of the orbital parameters of the NS. Increasing coherence time improves the sensitivity but increases the computational cost of the search. The values are chosen to roughly optimize the search (by maximizing the detection probability at fixed computing cost subject to some arbitrary assumptions about the prior on h_0) assuming a uniform prior on the projected semimajor axis $a \sin i$ and a Gaussian prior on the time of ascension T_{asc} . Longer coherence times are used for more likely values of T_{asc} (within $\pm 1\sigma$ of the mode of the prior distribution) and for smaller values of $a \sin i$ (where the parameter space metric of Whelan et al. 2015 implies a coarser resolution in T_{asc} and reduced computing cost).

associated with the large number of templates at different points in parameter space results in numerous candidates with $\rho \gtrsim 6$. A follow-up was performed whenever ρ exceeded a threshold of 6.5 for $25 \text{ Hz} < f_0 < 400 \text{ Hz}$, 6.2 for $400 \text{ Hz} < f_0 < 600 \text{ Hz}$, and 6.0 for $600 \text{ Hz} < f_0 < 2000 \text{ Hz}$. These thresholds were chosen in light of the number of templates searched (cf. Table 3) as a function of frequency. For each 5 Hz band, the threshold at which the expected number of Gaussian outliers was 0.1 (Figure 3). For simplicity, the three thresholds (6.5, 6.2, and 6.0) were chosen to be close to or slightly below these threshold values. As a result, the number of expected Gaussian outliers per 5 Hz was between 0.06 and 0.92. Table 3 shows the total expected number of outliers in each range of frequencies under the Gaussian assumption. Since the noise was not Gaussian, the actual number of signals followed up was significantly larger, as also shown in Table 3.

The follow-up procedure was as follows:

1. Data contaminated by known monochromatic noise features (“lines”) in each detector were excluded from the search from the start. In most cases, the time-dependent orbital Doppler modulation of the expected signal meant that a narrow line only affected data relevant to a subset of the SFTs from the run. Pairs involving these SFTs needed to be excluded from the sum in Equation (5) and the normalization in Equation (7). The impact of this is illustrated in Figure 6 (in Appendix A), which shows the reduction in the sensitivity ϑ from the omission of pairs from Equation (7).
2. Because a strong signal generally led to elevated statistic values over a range of frequencies, all of the candidates within 0.02 Hz of a local maximum were “clustered” together, with the location of the maximum determining the parameters of the candidate signal. These are known as the “level 0” results.
3. A “refinement” search was performed in a $13 \times 13 \times 13 \times 13$ grid in f_0 , with the same T_{\max} as the original

Table 2
Parameters Used for the Cross-correlation Search

| Parameter | Range |
|---------------------------------------|-------------------------------|
| f_0 (Hz) | [25, 2000] |
| $a \sin i$ (lt-s) ^a | [0.36, 3.25] |
| T_{asc} (GPS s) ^b | $1131415404 \pm 3 \times 179$ |
| P_{orb} (s) | $68023.70 \pm 3 \times 0.04$ |

Notes. Ranges for T_{asc} and P_{orb} are chosen to cover $\pm 3\sigma$ of the observational uncertainties, as illustrated in Figure 1.

^a The range for the projected semimajor axis, $a \sin i = K_1 P_{\text{orb}} / (2\pi)$, in light-seconds was taken from the constraint $K_1 \in [10, 90] \text{ km s}^{-1}$, which was the preliminary finding of Wang (2017) available at the time the search was constructed. Note that this range of $a \sin i$ values is broader than that used in previous analyses, which assumed a value from Abbott et al. (2007a) of 1.44 lt-s with a 1σ uncertainty of 0.18 lt-s.

^b This value for the time of ascension has been propagated forward by 3435 orbits from the value in Table 1, and corresponds to a time of 2015 November 13 02:03:07 UTC, near the middle of the O1 run. (This is useful when constructing the lattice to search over the orbital parameter space, as noted in Whelan et al. 2015.) The increase in uncertainty is due to the uncertainty in P_{orb} .

search, and $a \sin i$, T_{asc} , and P_{orb} centered on the original candidate, with a grid spacing chosen to be one-third of the original spacing (with appropriate modifications for P_{orb} depending on whether that parameter was resolved in the original search). This procedure produces a grid that covers ± 2 grid spacings of the original grid and has a mismatch of approximately $25\% \times (1/3)^2 \approx 2.8\%$. The results of this refinement stage are known as “level 1.”

4. A deeper follow-up was done on the level 1 results, with T_{\max} increased to $4 \times$ its original value. According to the theoretical expectation in Equation (7), this should approximately double the statistic value ρ for a true signal. Since this increase in coherence time also produces a finer parameter space resolution, the density of the grid was again increased by a further factor of 3 in each direction (resulting in a mismatch of approximately $25\% \times (1/3)^2 \times (4/3)^2 \approx 4.9\%$),¹⁵⁴ and the size of the grid was $13 \times 13 \times 13 \times 13$. The results of this follow-up stage are known as “level 2.” Signals whose detection statistic ρ decreases at this stage are dropped from the follow-up.
5. Surviving level 2 results were followed up by once again quadrupling the coherence time T_{\max} to $16 \times$ the original value, and increasing the density by a factor of 3 in each direction, for an approximate mismatch of $25\% \times (1/3)^2 \times (4/3)^2 \times (4/3)^2 \approx 8.8\%$. Again, true signals are expected to approximately double their statistic values, and the grid is modified as at level 2. The results of this round of follow-up are known as “level 3.”
6. Unknown instrumental lines in a single detector are likely to produce strong correlations between SFTs from that detector. To check for this, at each stage of follow-up, level 1 and beyond, the CrossCorr statistic ρ_{HH} was calculated using only data from LIGO Hanford Observatory (LHO), and the statistic ρ_{LL} using only data from LIGO Livingston Observatory (LLO). If we write ρ_{HL} as

¹⁵⁴ Note that the increased mismatch means that the highest S/N may not quite double, even for a true signal. As Figure 4 shows, simulated signals still show significant increases in S/N at levels 2 and 3 of the follow-up.

Table 3
Summary of Numbers of Templates and Candidates

| Min f_0 (Hz) | Max f_0 (Hz) | Min T_{\max} (s) | Max T_{\max} (s) | ρ Threshold | Number of Templates | Expected Gauss False Alarms ^a | Level 0 ^b | Level 1 ^c | Level 2 ^d | Level 3 ^e |
|----------------|----------------|--------------------|--------------------|------------------|-----------------------|--|----------------------|----------------------|----------------------|----------------------|
| 25 | 50 | 10,080 | 25,920 | 6.5 | 1.58×10^{10} | 0.6 | 269 | 212 | 62 | 6 |
| 50 | 100 | 8160 | 19,380 | 6.5 | 7.96×10^{10} | 3.2 | 499 | 473 | 209 | 14 |
| 100 | 150 | 6720 | 15,120 | 6.5 | 1.51×10^{11} | 6.1 | 605 | 571 | 304 | 29 |
| 150 | 200 | 5040 | 11,520 | 6.5 | 1.62×10^{11} | 6.5 | 456 | 432 | 260 | 35 |
| 200 | 300 | 2400 | 6600 | 6.5 | 1.33×10^{11} | 5.3 | 220 | 194 | 87 | 29 |
| 300 | 400 | 1530 | 4080 | 6.5 | 6.62×10^{10} | 2.7 | 254 | 216 | 23 | 10 |
| 400 | 600 | 360 | 1800 | 6.5 | 1.62×10^{10} | 0.6 | 88 | 26 | 2 | 1 |
| 600 | 800 | 360 | 720 | 6.2 | 5.80×10^9 | 1.6 | 78 | 15 | 2 | 2 |
| 800 | 1200 | 300 | 300 | 6.0 | 1.18×10^{10} | 11.7 | 145 | 134 | 3 | 0 |
| 1200 | 2000 | 240 | 240 | 6.0 | 3.12×10^{10} | 30.8 | 442 | 107 | 6 | 1 |

Notes. For each range of frequencies, this table shows the minimum and maximum coherence time T_{\max} used for the search, across the different orbital parameter space cells (see Figure 2), the threshold in signal-to-noise ratio (S/N) ρ used for follow-up, the total number of templates, and the number of candidates at various stages of the process. (See Section 4 for detailed description of the follow-up procedure.)

^a This is the number of candidates that would be expected in Gaussian noise, given the number of templates and the follow-up threshold.

^b This is actual number of candidates (after clustering) that crossed the S/N threshold and were followed up.

^c This is the number of candidates remaining after refinement. All of the candidates “missing” at this stage have been removed by the single-detector veto for unknown lines.

^d This is the number of candidates remaining after each one has been followed up with a T_{\max} equal to $4 \times$ the original T_{\max} for that candidate. (True signals should approximately double their S/N; any candidates whose S/N goes down have been dropped.) All of the signals present at this stage are shown in Figure 4, which also shows the behavior of the search on simulated signals injected in software.

^e This is the number of candidates remaining after T_{\max} has been increased to $16 \times$ its original value.

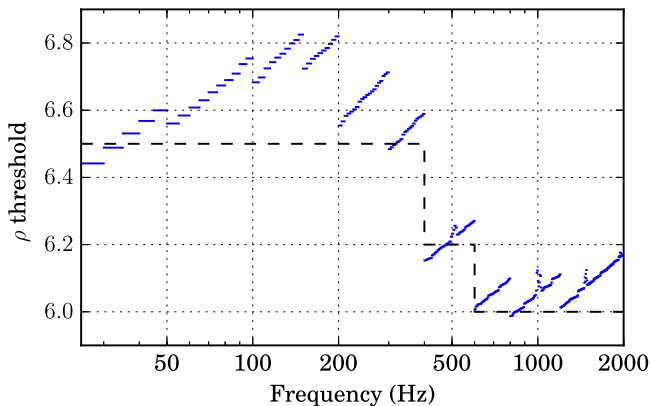


Figure 3. Selection of follow-up threshold as a function of frequency. If the data contained no signal and only Gaussian noise, each template in parameter space would have some chance of producing a statistic value exceeding a given threshold. Within each 0.05 Hz frequency band, the total number of templates was computed and used to find the threshold at which the expected number of Gaussian outliers above that value would be 0.1 (short blue lines). For simplicity, the actual follow-up threshold was chosen near or below that level, producing thresholds of 6.5 for $25 \text{ Hz} < f_0 < 400 \text{ Hz}$, 6.2 for $400 \text{ Hz} < f_0 < 600 \text{ Hz}$, and 6.0 for $600 \text{ Hz} < f_0 < 2000 \text{ Hz}$ (black dashed line). Note that the large number of non-Gaussian outliers (cf. Table 3) makes the Gaussian follow-up level an imprecise tool in any event.

the statistic constructed using only pairs of one SFT from LHO and one from LLO, the overall statistic can be written (cf. Equations (2.36), (3.6) and (3.7) of Whelan et al. 2015) as

$$\rho = \frac{\vartheta_{\text{HH}}\rho_{\text{HH}} + \vartheta_{\text{LL}}\rho_{\text{LL}} + \vartheta_{\text{HL}}\rho_{\text{HL}}}{\vartheta}, \quad (8)$$

where

$$\vartheta = \sqrt{\vartheta_{\text{HH}}^2 + \vartheta_{\text{LL}}^2 + \vartheta_{\text{HL}}^2}. \quad (9)$$

Since, for example, $E[\rho] = (h_0^{\text{eff}})^2 \vartheta > (h_0^{\text{eff}})^2 \vartheta_{\text{HH}} = E[\rho_{\text{HH}}]$, we expect true signals to have higher overall detection statistics ρ than the single-detector statistics ρ_{HH} and ρ_{LL} . We therefore veto any candidate for which $\rho_{\text{HH}} > \rho$ or $\rho_{\text{LL}} > \rho$ at any level of follow-up. This is responsible for the reduction of candidates from level 0 to level 1 seen in Table 3.

A total of 127 candidates survive level 3 of the follow-up. To check whether any of them represent convincing detection candidates, we plot in Figure 4 the ratio by which the S/N increases from level 1 to level 2, and from level 2 to level 3. We also plot the corresponding ratios for all of the candidates surviving level 2 (the $16 \times$ original T_{\max} follow-up is not available for candidates that fail level 2), and also for the simulated signal injections described in Appendix A. We see that none of the candidates come close to doubling their S/N at either stage; in fact, none of them even double their S/N from level 1 to level 3. We empirically assess the follow-up procedure with the injections, and find that their S/Ns generally increase by slightly less than the naïvely expected factor of 2 (perhaps because of the increasing mismatch at later follow-up levels). We do see that the injected signals (at least those that survive level 2 follow-up and appear on the plot) nearly all increase their S/N noticeably more than any of the candidates from the search. Also note that of the 666 injected signals (out of 754) that produced ρ values above their respective thresholds, 652 survived all levels of follow-up. (There were four vetoed at level 1, four at level 2, and six at level 3 of the follow-up.) All but a handful of those 652 (between one and four, depending on the stringency of the criterion) are well-separated from the bulk of the search results in Figure 4. We thus conclude that our follow-up procedure is relatively robust, and that there are no convincing detection candidates from the search.

The signal model in this search assumes that the GW frequency f_0 in the NS’s reference frame is constant. In practice, the

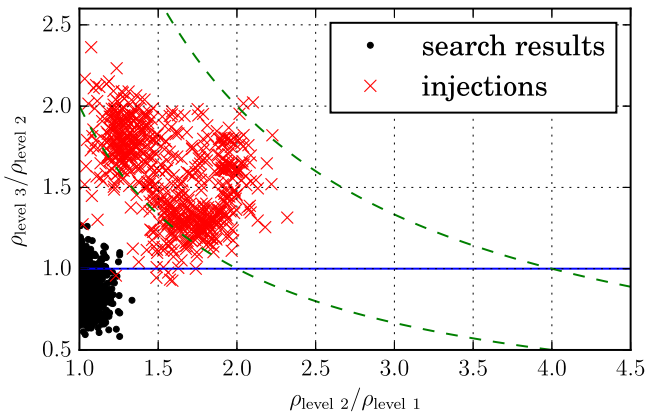


Figure 4. Ratios of follow-up statistics for search candidates and simulated signals. This plot shows all of the candidates that survived to level 2 of the follow-up (see Section 4 and Table 3), both from the main search and from the analysis of the simulated signal injections described in Appendix A. It shows the ratios of the S/N after follow-up level 1 (at the original coherence time T_{max}), level 2 (at $4\times$ the original coherence time), and level 3 (at $16\times$ the original coherence time). The green dashed lines are at constant values of $\rho_{\text{level } 3} / \rho_{\text{level } 1}$ equal to 2 and 4, respectively. There are no points with $\rho_{\text{level } 2} / \rho_{\text{level } 1} < 1$, because those candidates do not survive level 2 follow-up and are therefore not subjected to level 3 follow-up. From Equation (5) and Equation (7), the naïve expectation is that the S/N will roughly double each time T_{max} is quadrupled. Empirically, the follow-ups of injections do not show exactly that relationship, but the vast majority do show significant increases in S/N, which are not seen in *any* of the follow-ups of search candidates, leading to the conclusion that no convincing detection candidates are present.

equilibrium in an LMXB will be only approximate, and the intrinsic frequency will vary stochastically with time. Whelan et al. (2015) estimated the effect of spin wandering under a simplistic random-walk model in which the GW frequency underwent a net spin-up or spin-down of magnitude $|\dot{f}|_{\text{drift}}$, changing on a timescale T_{drift} . The fractional loss of S/N was estimated as

$$\frac{E[\rho]_{\text{ideal}} - E[\rho]}{E[\rho]_{\text{ideal}}} \approx \frac{\pi^2}{6} T_{\text{run}} T_{\text{drift}} |\dot{f}|_{\text{drift}}^2 T_{\text{max}}^2, \quad (10)$$

where T_{run} is the duration of the observing run from the start to end, not considering duty factors (in contrast to the T_{obs} appearing in Equation (7)) or numbers of detectors. To give an illustration of the possible impacts of spin wandering on the present search, we make reference to the values of $|\dot{f}|_{\text{drift}} = 10^{-12} \text{ Hz s}^{-1}$ and $T_{\text{drift}} = 10^6 \text{ s}$. These are conservative upper limits on how fast the signal can drift, based on Bildsten (1998). Similar values have been used in the first Sco X-1 mock data challenge (Messenger et al. 2015) and other work on Sco X-1 (Leaci & Prix 2015; Whelan et al. 2015).¹⁵⁵

In the O1 run, where the run duration was $T_{\text{run}} = 1.12 \times 10^7 \text{ s}$, the theoretical fractional loss of S/N will be

$$0.012 \left(\frac{T_{\text{drift}}}{10^6 \text{ s}} \right) \left(\frac{|\dot{f}|_{\text{drift}}}{10^{-12} \text{ Hz s}^{-1}} \right)^2 \left(\frac{T_{\text{max}}}{25000 \text{ s}} \right)^2. \quad (11)$$

Since our largest initial T_{max} value is 25,290 s, the impact on the initial search and the upper limit of spin wandering at or below this level would be negligible. Note that even spin wandering,

¹⁵⁵ For comparison, the maximum spin wandering that could be tracked by the Viterbi analysis of Abbott et al. (2017f) is $|\dot{f}|_{\text{drift}} = (\Delta f_{\text{drift}}) / T_{\text{drift}} = 1 / (2T_{\text{drift}}^2) \approx 0.7 \times 10^{-12} \text{ Hz s}^{-1}$ at $T_{\text{drift}} = 10 \text{ d} \approx 0.9 \times 10^6 \text{ s}$.

which posed no complication for the initial search, could potentially be a limitation for the follow-up procedure, where T_{max} is increased by a factor of 4 at level 2 and a factor of 16 at level 3. In any event, the impact depends on the level of spin wandering present, which is still an area of open research.

5. Upper Limits

In the absence of a detection, we set upper limits on the strength of gravitational radiation from Sco X-1, as a function of frequency. We used as a detection statistic ρ_{max} , the maximum statistic value observed in a 0.05 Hz band. We produced frequentist 95% upper limits via a combination of theoretical considerations and calibration with simulated signals, as explained in detail in Appendix A. The starting point was a Bayesian upper limit constructed using the expected statistical properties of the detection statistic and corrected for the reduction of sensitivity due to known lines. A series of simulated signal injections was then performed and used to estimate a global adjustment factor to estimate the amplitude at which a signal would have a 95% chance of increasing the ρ_{max} value in a band.

The procedure produced two sets of upper limits: a limit on h_0 including marginalization over the unknown inclination angle ι , and an unmarginalized limit on the quantity h_0^{eff} defined in Equation (2) to which the search is directly sensitive. The h_0^{eff} upper limit can also be interpreted as a limit on h_0 subject to the assumption of circular polarization (optimal spin orientation corresponding to $\cos \iota = \pm 1$). It can be converted to a limit assuming linear polarization $\cos \iota = 0$ by multiplying by $\sqrt{8} = 2.83$. If we assume that the NS spin is aligned with the binary orbit (as one would expect for an NS spun up by accretion), $\iota \approx i \approx 44^\circ$, we obtain a limit on h_0 , which is the h_0^{eff} upper limit multiplied by 1.35.

We show the marginalized and unmarginalized upper limits of this search in Figure 5, along with the other upper limits on Sco X-1 set with O1 data: the unmodeled stochastic radiometer (Ballmer 2006) results of Abbott et al. (2017d) and the directed search results of Abbott et al. (2017f) using Viterbi tracking of a hidden Markov model (Suvorova et al. 2016) to expand the applicability of the sideband search (Messenger & Woan 2007; Sammut et al. 2014; Aasi et al. 2015b) over the whole run. The present results improve on these by a factor of 3–4, yielding a marginalized limit of $h_0 \lesssim 2.3 \times 10^{-25}$ and an unmarginalized limit of $h_0^{\text{eff}} \lesssim 8.0 \times 10^{-26}$ at the most sensitive signal frequencies between around 100 Hz and 200 Hz. The marginalized 95% upper limits from Initial LIGO data (Abadie et al. 2011; Aasi et al. 2015b; Meadors et al. 2017) were all around 1.5×10^{-25} , so we have achieved an overall improvement of a factor of 6–7 from Initial LIGO to Advanced LIGO’s first observing run, a combination of decreased detector noise and algorithmic improvements.

We also plot for comparison the torque balance level predicted by Equation (4). The marginalized limits on h_0 come closest to this level at 100 Hz, where they are within a factor of 3.4 of this theoretical level. In terms of h_0^{eff} , the torque balance level depends on the unknown value of the inclination ι . For the most optimistic case of circular polarization ($\cos \iota = \pm 1$), our unmarginalized limit is a factor of 1.2 above the torque balance level near 100 Hz. Assuming linear polarization puts our limits within a factor of 3.5 of this level, and the most likely

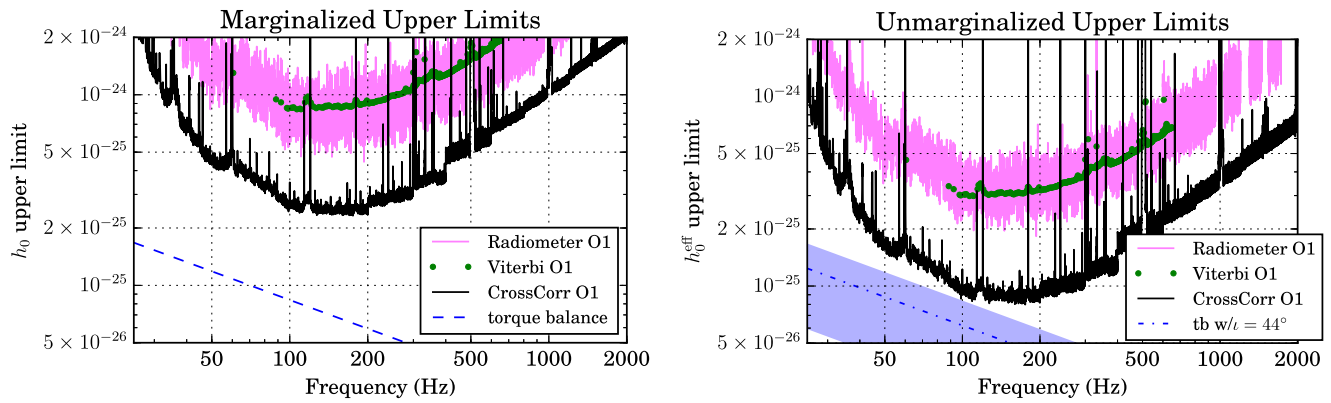


Figure 5. Upper limits from directed searches in O1 data. Left: upper limit on h_0 , after marginalizing over the neutron star spin inclination ι , assuming an isotropic prior. The dashed line shows the nominal expected level assuming torque balance (Equation (4)) as a function of frequency. Right: upper limit on h_0^{eff} , defined in Equation (2). This is equivalent to the upper limit on h_0 assuming circular polarization. (Note that the marginalized upper limit on the left is dominated by linear polarization, and so is a factor of $\sim\sqrt{8}$ higher.) The shaded band shows the range of h_0^{eff} levels corresponding to the torque balance h_0 in the plot on the left, with circular polarization at the top and linear polarization on the bottom. The dotted–dashed line (labelled “tb w/ $\iota = 44^\circ$ ”) corresponds to the assumption that the neutron star spin is aligned to the most likely orbital angular momentum, and $\iota \approx i \approx 44^\circ$. (See Table 1.) For comparison with the “CrossCorr” results presented in this paper, we show the “unknown polarization” and “circular polarization” curves from the Viterbi analysis in Abbott et al. (2017f; dark green dots), as well as the 95% marginalized and circular polarization adapted from the Radiometer analysis in Abbott et al. (2017d; broad light magenta curve). Note that the Viterbi analysis reported upper limits for the 1 Hz bands, while the current CrossCorr analysis does so for the 0.05 Hz bands, and the Radiometer analysis for the 0.03125 Hz bands. This gives the upper limit curves for CrossCorr and especially Radiometer a “fuzziness” associated with noise fluctuations between adjacent frequencies rather than any physically meaningful distinction. When comparing 95% upper limits between the different analyses, it is therefore appropriate to look near the 95th percentile of this “fuzz” rather than at its bottom.

value of $\iota = 44$ corresponds to an upper limit curve a factor of 1.7 above the torque balance level, again near 100 Hz.

6. Outlook for Future Observations

We have presented the results of a search for GWs from Sco X-1 using data from Advanced LIGO’s first observing run. The upper limits on the GW amplitude represent a significant improvement over the results from Initial LIGO and are within a factor of 1.2–3.5 of the benchmark set by the torque balance model, depending on assumptions about system orientation. Future observing runs (Abbott et al. 2016f) are expected to produce an improvement in the detector strain sensitivity of $\gtrsim 2.5$. An additional enhancement will come with longer runs, as the amplitude sensitivity of the search scales as $T_{\text{obs}}^{1/4}$. Algorithmic improvements that allow larger T_{max} with the same computing resources will also lead to improvements, as the sensitivity scales as $T_{\text{max}}^{1/4}$ as well. A promising area for such an improvement is the use of resampling (Patel et al. 2010) to reduce the scaling of computing cost with T_{max} (G. D. Meadors et al. 2017, in preparation). (A similar method is used in the proposed semicoherent search described in Leaci & Prix 2015.) These anticipated instrumental and algorithmic improvements make it likely that search sensitivities will surpass the torque balance level over a range of frequencies (as projected in Whelan et al. 2015), and suggest the possibility of a detection during the advanced detector era, depending on details of the system such as GW frequency, inclination of the NS spin to the line of sight, and how close the system is to GW torque balance.

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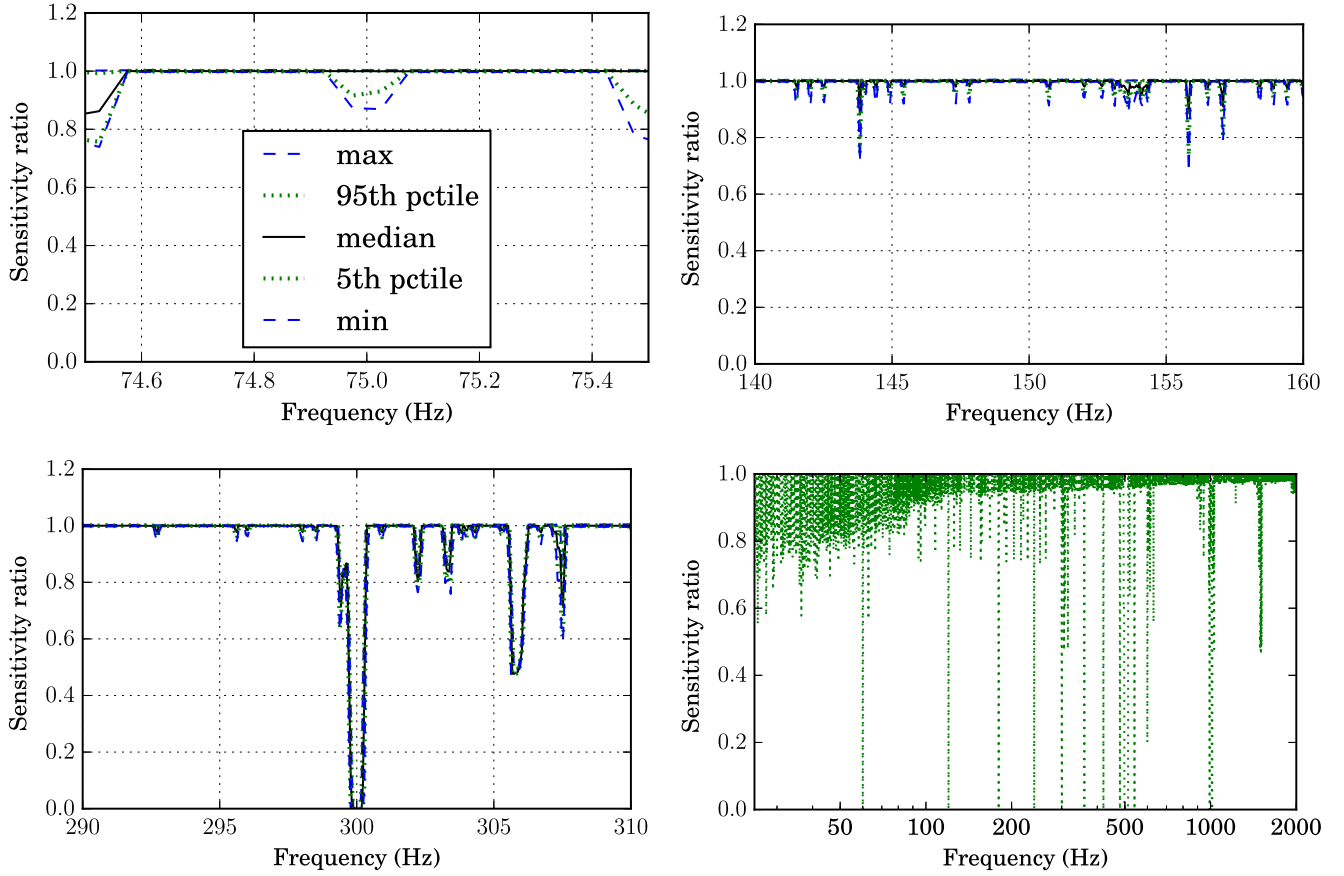


Figure 6. Impact of known lines on the sensitivity of the search. Fourier bins impacted by known lines are removed from the calculation of the statistic ρ defined in Equation (5) and from the sensitivity $\vartheta = E[\rho]/(h_0^{\text{eff}})^2$ defined in Equation (7). For a given signal frequency f_0 , data are removed at some times due to the time-varying Doppler shift which depends on the orbital parameter $a \sin i$. The effect is to lower ϑ relative to the value it would have if the lines were not removed; this “sensitivity ratio” goes to zero if all of the data relevant to a signal frequency f_0 are removed by the line. The first three plots contain illustrations of the percentiles of this ratio, taken over intervals of width 0.05 Hz. (There is a range of values in each frequency interval because of its finite width, and the range of $a \sin i$ values which determine the magnitude of the Doppler modulation.) Note that the broad line at 300 Hz (a harmonic of the 60 Hz AC power line) effectively nullifies the search at that frequency. The last plot shows the fifth percentile of the sensitivity ratio in 0.05 Hz intervals across the whole sensitivity band.

Appendix A Details of the Upper Limit Method

The method used to set the upper limits for each 0.05 Hz band in Section 5 consisted of three steps:

1. An idealized 95% Bayesian upper limit was constructed using the posterior distribution $\text{pdf}(h_0^2 | \rho_{\text{max}})$ or $\text{pdf}([h_0^{\text{eff}}]^2 | \rho_{\text{max}})$.
2. A correction factor was applied in each 0.05 Hz band to account for the loss of sensitivity due to omission of data impacted by known lines.
3. A series of software injections was performed near the level of the 95% upper limit and used to empirically estimate a global correction factor for each upper limit curve based on the recovery or non-recovery of the injections.

A.1. Idealized Bayesian Method

The Bayesian calculation assumes that all of the ρ values for templates in the initial search represent independent Gaussian random variables with unit variance; one has mean $[h_0^{\text{eff}}]^2 \vartheta$ and the others have zero mean. Note that different regions of the orbital parameter space have different coherence times T_{max} and therefore ϑ values (cf. Equation (7)). The method produces a

sampling distribution $\text{pdf}(\rho_{\text{max}} | [h_0^{\text{eff}}]^2)$, marginalizing over the location of the signal in orbital parameter space.

This sampling distribution is used to construct a posterior distribution $\text{pdf}([h_0^{\text{eff}}]^2 | \rho_{\text{max}})$ assuming a uniform prior in $(h_0^{\text{eff}})^2$, and this is used to produce a 95% Bayesian upper limit on $(h_0^{\text{eff}})^2$ according to

$$\int_0^{(h_0^{\text{eff}})_{95\%}^2} d[h_0^{\text{eff}}]^2 \text{pdf}([h_0^{\text{eff}}]^2 | \rho_{\text{max}}) = 0.95. \quad (12)$$

To produce an upper limit on the intrinsic strength h_0 , we assume a prior that is uniform in h_0^2 and $\cos \iota$, repeat the calculation above, and numerically marginalize over $\cos \iota$ to obtain a posterior $\text{pdf}(h_0^2 | \rho_{\text{max}})$.

A.2. Correction for Known Lines

Although we calculate a single ϑ value for each of the 18 search regions for a given 0.05 Hz band and use it in the calculation, the search can in principle have a different ϑ value for each template. This is because of the correction which omits data contaminated by Doppler-modulated known instrumental lines from the sum in Equation (7), a process that depends on the signal frequency f_0 as well as the projected orbital semimajor

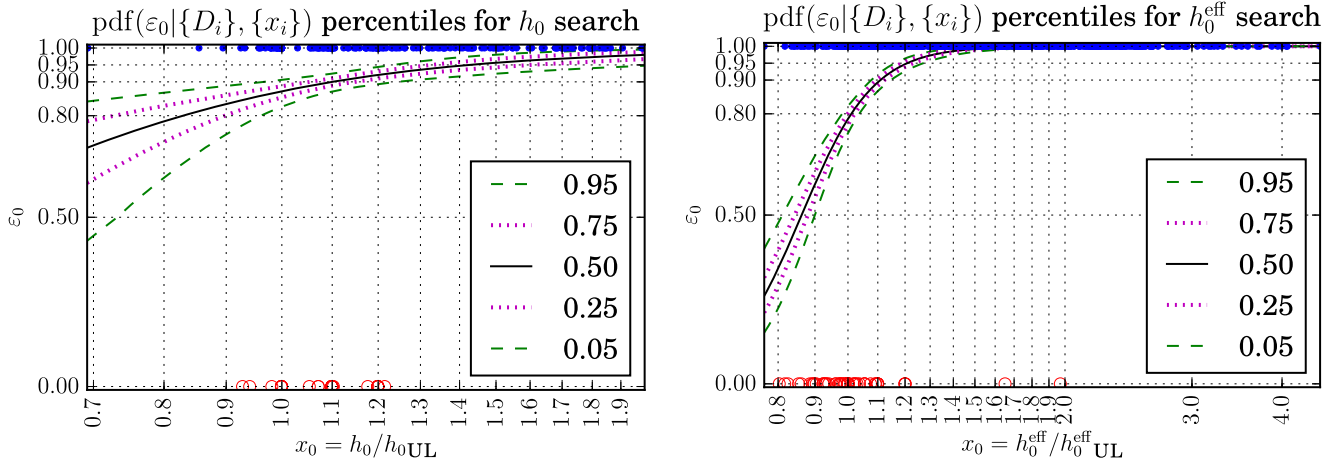


Figure 7. Estimation of efficiency from recovery of simulated signals injected in the software. At left, the results of the 376 injections with amplitude h_0 specified in terms of the uncorrected marginalized upper limit $h_{0\text{UL}}$ are shown as black dots, with recovered injections (those that increased the maximum S/N ρ_{max} in the relevant 0.05 Hz band) shown as blue dots on the $\varepsilon_0 = 1$ line and unrecovered injections shown as red circles at $\varepsilon_0 = 0$. The recovered and unrecovered injections are used to produce a posterior $\text{pdf}(\alpha, \beta|\{D_i\}, \{x_i\})$ according to Equation (13), and this is used to generate posterior distributions $\text{pdf}(\varepsilon_0|\{D_i\}, \{x_i\})$ at a range of signal strengths $x_0 = h_0/h_{0\text{UL}}$ according to Equation (14); the 5th, 25th, 50th, 75th, and 95th percentiles are calculated from these distributions as a function of x_0 . The x_0 value at which the posterior expectation $E(\varepsilon_0|\{D_i\}, \{x_i\})$ crosses 0.95 is used as a correction factor by which we multiply $h_{0\text{UL}}$ to produce the final marginalized upper limit shown in Figure 5. At right, we do the same thing for h_0^{eff} , using the full set of 754, and derive a correction factor by which to multiply the unmarginalized upper limit $h_{0\text{UL}}^{\text{eff}}$. Note that in the h_0 search the value of x_0 corresponding to $\varepsilon_0 = 0.95$ is less accurately determined than that for the h_0^{eff} search, both because of the smaller number of applicable injections and because the detection efficiency depends more weakly on h_0 .

axis $a \sin i$. In each 0.05 Hz band, we estimate the distribution of the ratio of the actual ϑ to the band-wide ϑ ; the percentiles of this distribution are illustrated in Figure 6. We divide by the fifth percentile of this distribution (shown in the last panel of Figure 6) to produce corrected h_0^2 and $(h_0^{\text{eff}})^2$ upper limits.

A.3. Empirical Adjustment from Software Injections

We performed a series of re-analyses of the data with a total of 754 simulated signals (“software injections”) added to the data stream to validate the upper limits including the known line correction. The signals were generated over signal frequencies from 25 to 500 Hz, some with h_0 set to some multiple of the marginalized 95% upper limit $h_{0\text{UL}}$, and others with h_0^{eff} set to some multiple of the unmarginalized 95% upper limit $h_{0\text{UL}}^{\text{eff}}$. We defined “recovery” of the injection as an increase in the maximum detection statistic ρ_{max} compared to the results with no signal present. (Follow-up of the recovered injections that crossed the relevant ρ threshold was also performed as a way of testing our follow-up procedure, as described in Section 4.) We find that the fraction of signals of each type recovered when the injection is done at the upper limit level to be slightly below the expected 95%.¹⁵⁶ This is to be expected, as there are various approximations in the method, such as the tolerated mismatch in the initial parameter space grid and the acceptable loss of S/N due to finite-length SFTs, which should lead to an S/N slightly less than that predicted by Equation (6).

To estimate empirically the amount by which the upper limits should be scaled to produce a 95% injection recovery efficiency, we apply the method described in Whelan (2015) and used to produce the efficiency curves in Messenger et al. (2015). We posit a simple sigmoid model where the efficiency

of the search as a function of signal strength x is assumed to be $\varepsilon(x; \alpha, \beta) = (1 + e^{-\alpha(\ln x - \beta)})^{-1}$ and construct the posterior from the recovery data ($D_i = 1$ if the signal i was recovered, 0 if not):

$$\text{pdf}(\alpha, \beta|\{D_i\}, \{x_i\}) \propto \prod_i \varepsilon(x_i; \alpha, \beta)^{D_i} \times (1 - \varepsilon(x_i; \alpha, \beta))^{1-D_i} \text{pdf}(\alpha, \beta). \quad (13)$$

With sufficient data, the prior should be irrelevant, but we take a noninformative prior $\text{pdf}(\alpha, \beta) \propto \alpha^{-1}$ and define the signal strength x as the h_0 or h_0^{eff} of the injection divided by the corresponding upper limit. We can then construct, at any signal level x_0 , the posterior on the efficiency $\varepsilon_0 = \varepsilon(x_0; \alpha, \beta)$, marginalized over α and β :

$$\text{pdf}(\varepsilon_0|\{D_i\}, \{x_i\}) = \int_0^\infty d\alpha \int_{-\infty}^\infty d\beta \text{pdf}(\alpha, \beta|\{D_i\}, \{x_i\}) \delta(\varepsilon_0 - \varepsilon(x_0; \alpha, \beta)). \quad (14)$$

The posterior distributions of efficiency are shown in Figure 7. We define the correction factor to be the x_0 at which the expectation value $\int_0^1 d\varepsilon_0 \text{pdf}(\varepsilon_0|\{D_i\}, \{x_i\})$ crosses 95%.

A total of eight sets of injections were performed, four with h_0 at a specified multiple of $h_{0\text{UL}}$, and four with h_0^{eff} at a specified multiple of $h_{0\text{UL}}^{\text{eff}}$. The multipliers were 1.0, 1.1, 1.2, and a random value between 1.1 and 2.0 chosen from a log-uniform distribution. For the unmarginalized h_0^{eff} upper limit, we use all eight sets of injections, 754 in total and find the expectation value of the efficiency crosses 95% at $h_0^{\text{eff}}/h_{0\text{UL}}^{\text{eff}} \approx 1.21$. This factor has been applied to $h_{0\text{UL}}^{\text{eff}}$ to produce the upper limits in Figure 5.

For the marginalized h_0 upper limit, we must restrict ourselves to the four injection sets which specified $h_0/h_{0\text{UL}}$. This is because our search is primarily sensitive to h_0^{eff} , and specifying h_0^{eff} while choosing the inclination angle i randomly implies anticorrelations between h_0 and $|\cos i|$. Signals with

¹⁵⁶ The fraction of signals recovered is a frequentist statement, as opposed to the Bayesian upper limit constructed from the posterior, but the two types of upper limits are related closely enough (see, for example, Rover et al. 2011) that the fraction should be close to 95%.

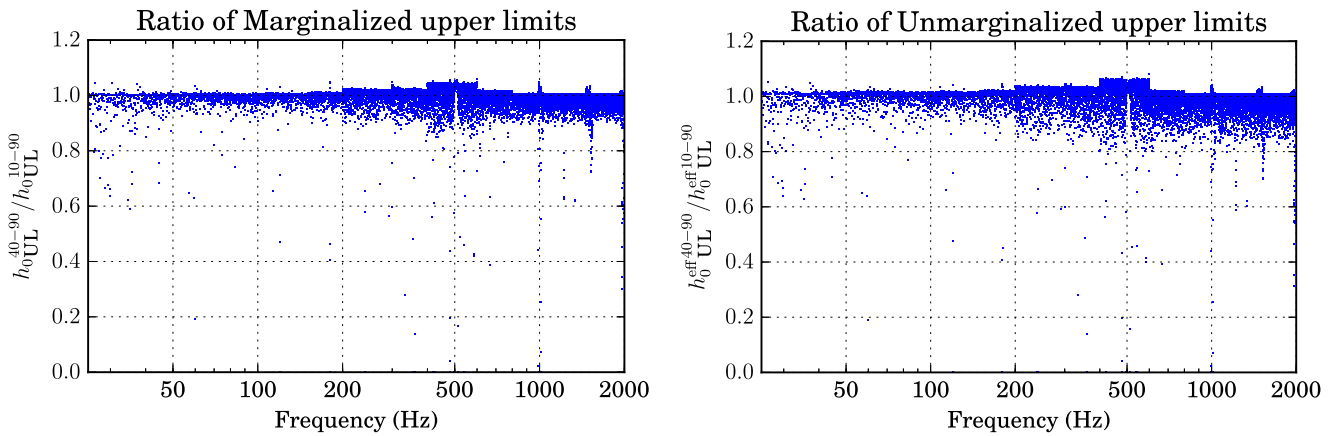


Figure 8. Comparison of upper limits constructed by restricting attention to $a \sin i \geq 1.44$ lt-s ($K_1 \geq 40$ km s $^{-1}$) to those from the original search. The results are generally comparable; we plot the ratio of the upper limits rather than reproducing the curves in Figure 5, because the changes in the latter would barely be noticeable. The step-like features that are visible are due to the details of the search (such as T_{\max} values) being different in different frequency ranges listed in Table 3.

high h_0 values will tend to be those with unfavorable polarization, and therefore not be any easier to detect. Using the 376 applicable injections, we estimate the 95% efficiency at $h_0/h_{0UL} \approx 1.44$ and use this factor when generating the final upper limit shown in Figure 5. Note that this is less well determined than the factor for the unmarginalized h_0^{eff} upper limit. This is both because of the smaller number of injections used and because h_0 correlates less well with detectability than h_0^{eff} . However, the upper limit curve for h_0 is very close to the unmarginalized upper limit assuming linear polarization ($\cos \iota = 0$), which is consistent with the expectation that the 95% upper limit will be dominated by this worse-case scenario.

Appendix B

Results with a Constrained Semimajor Axis

As noted in Table 2, the range of $a \sin i$ values searched was chosen based on preliminary information from Wang (2017), which constrained the projected orbital velocity K_1 to lie between 10 and 90 km s $^{-1}$. This was subsequently refined to between 40 and 90 km s $^{-1}$. For comparison, we recomputed the upper limits, discarding the results of searches with $a \sin i \leq 1.44$ lt-s, corresponding to the nine bottom-most parameter space cells shown in Figure 2. The results were not significantly different (for instance, they were barely noticeable on plots like Figure 5), but for illustration we plot in Figure 8 the ratio of the two sets of upper limits. A bigger impact of the refined parameter space will be seen in future runs, when computing resources can be concentrated on the allowed range of $a \sin i$ values.

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